Fractional snow-covered area: Scale-independent peak of winter parameterization

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Abstract. The spatial distribution of snow in the mountains is significantly influenced through interactions of topography with wind, precipitation, shortwave and longwave radiation, and avalanches that may relocate the accumulated snow. One of the most crucial model parameters for various applications such as weather forecasts, climate predictions and in hydrological modeling is the fraction of the ground surface that is covered by snow, also called fractional snow-covered area ($fSCA$). While previous subgrid parameterizations for the spatial snow depth distribution and $fSCA$ work well, performances were scale-dependent. Here, we were able to confirm a previously established empirical relationship of the peak of winter parameterization for the standard deviation of snow depth σ_{HS} by evaluating it on 11 spatial snow depth data sets from 7 different geographic regions and snow climates with resolutions ranging from 0.1 m to 3 m. Enhanced performance (mean percentage errors (MPE) decreased by 25 %) across all spatial scales ≥ 200 m was achieved by recalibrating and introducing a scale-dependency in the dominant scaling variables. Scale-dependent MPEs vary between -7 % and 3 % for σ_{HS} and between 0 % and 1 % for $fSCA$. We performed a scale- as well as region-dependent evaluation of the parameterizations to assess the potential performances on independent data sets. This evaluation revealed that for the majority of the regions the MPEs mostly lie between ± 10 % for σ_{HS} and between -1 % and 1.5 % for $fSCA$. This suggests that the new parameterizations perform similarly well in most geographical regions.

1 Introduction

Whenever there is snow on the ground, there will be large spatial variability in snow depth. In mountainous terrain, this spatial distribution of snow is significantly influenced by topography due to corresponding spatial variations in wind, precipitation, shortwave and longwave radiation, and in steep terrain by avalanches that may relocate the accumulated snow. As a result, the snow-covered landscape can consist of a complex mix of snow-free and snow-covered areas, especially in steep terrain or during snow melt. A parameter which describes how much of the ground is covered by snow is the fractional snow-covered area ($fSCA$). Most of the time $fSCA$ is tightly linked to snow depth (HS) and in particular to its spatial distribution. $fSCA$

is able to bridge between the spatial mean HS and the actual observed snow coverage. Sound $fSCA$ models are therefore crucial, since for the same spatial mean HS in early winter and in late spring the associated $fSCA$ can be completely different (e.g. Luce et al., 1999; Niu and Yang, 2007; Magand et al., 2014).

25    $fSCA$ plays a key role in modelling physical processes for various applications such as weather forecasts (e.g. Douville et al., 1995; Doms et al., 2011), climate simulations (e.g. Roesch et al., 2001; Mudryk et al., 2020) and avalanche forecasting (Bellaire and Jamieson, 2013; Horton and Jamieson, 2016; Vionnet et al., 2016). As climate warms, $fSCA$ is an highly relevant indicator for spatial snow cover changes in climate projections (e.g. Mudryk et al., 2020). A decrease in spatial snow extent prominently changes surface characteristics such as albedo in mountain landscapes, leading to changes in surface radiation, a

30   primary component of the surface energy balance. $fSCA$ is also a parameter in hydrological models to scale water discharges ~~in the different model grid cells managing in this way appropriately~~ appropriately to help manage basins water supply (e.g. Luce et al., 1999; Thirel et al., 2013; Magnusson et al., 2014; Griessinger et al., 2016). Errors in $fSCA$ estimates directly translate into errors of snow melt rates and melt water discharge (Magand et al., 2014). Thus, accurately describing $fSCA$ is of key importance for multiple model applications in mountainous terrain where highly variable spatial snow distributions

35   occur.

A parameterization of $fSCA$ can be obtained from satellite remote sensing products using optical imagery with varying spatiotemporal resolution. For instance, Sentinel-2 gathers data at a spatial resolution of 10 to 20 m at frequent global revisit intervals (<5 days, cloud-permitting) (Drusch et al., 2012; Gascoin et al., 2019). The availability of satellite-derived $fSCA$ remains however inconsistent due to time gaps between satellite revisits, data delivery and the frequent presence of clouds, which obscure the ground,

40   especially in winter in mountainous terrain reducing the availability of images drastically (e.g. Parajka and Blöschl, 2006; Gascoin et al., 2015). Satellite-derived $fSCA$ can also not be used directly for forecasting. Alternatively, $fSCA$ can be obtained from spatially averaging by using snow models at subgrid scales. While such snow cover models are available (e.g. Tarboton et al., 1996; Marks et al., 1999; Lehning et al., 2006; Essery et al., 2013; Vionnet et al., 2016), up until now they cannot be used in very high spatial resolutions over very large regions, in part due to a lack of detailed input data, such as

45   fine-scale surface wind speed and precipitation, as well as due to high computational cost. Often they are limited by model parameters and structure requiring calibration. Integrating data assimilation algorithms in snow models is able to mitigate some of these limitations which led for instance to improvements in runoff simulations (e.g. Andreadis and Lettenmaier, 2006; Nagler et al., 2008; Thirel et al., 2013; Griessinger et al., 2016; Huang et al., 2017; Griessinger et al., 2019). However, uncertainties inherently present in the input or assimilation data still remain, which are generally accentuated over snow-

50   covered catchments (Raleigh et al., 2015). Today, $fSCA$ parameterizations describing the subgrid snow depth variability therefore remain unavoidable for complex model systems and to complement assimilation of satellite-retrieved $fSCA$ products especially over mountainous terrain.

A parameterization of $fSCA$ describes the relationship between $fSCA$ and grid cell-averaged HS or snow water equivalent (SWE) by a so-called snow-cover depletion (SCD) curve. SCD curves were originally introduced in models without

55   taking into account subgrid topography or vegetation. In principle, there are two commonly applied forms: so-called closed functional forms and parametric probabilistic SCD curve formulations (Essery and Pomeroy, 2004). Parametric SCD curves

have disadvantages for practical applications such as numerical stability, computational efficiency and assuming an unimodal distribution which might be less appropriate for large grid cells covering heterogeneous surface such as mountainous terrain (e.g. Essery and Pomeroy, 2004; Swenson and Lawrence, 2012). Various closed functional forms for $fSCA$ are therefore ap-

60   plied in land surface and climate models (e.g. Douville et al., 1995; Roesch et al., 2001; Yang et al., 1997; Niu and Yang, 2007; Su et al., 2008; Swenson and Lawrence, 2012). Most of these parameterizations use simple relationships between $fSCA$ and spatial mean HS or SWE. Since topography strongly determines the spatial snow depth or snow water equivalent distribution (Clark et al., 2011), in the past, terrain characteristics were mostly heuristically introduced in closed form curves to account for subgrid terrain influences on $fSCA$ (e.g. Douville et al., 1995; Roesch et al., 2001; Swenson and Lawrence, 2012). To verify

65   the commonly applied closed forms of $fSCA$, Essery and Pomeroy (2004) integrated over log-normal SWE distributions and fitted the parametric SCD curves. The best obtained fit resulted for a function proportional to $tanh$ which is a previously derived closed form from Yang et al. (1997). By using a normal probability density function (pdf) Helbig et al. (2015) obtained the same form fit for $fSCA$ as Essery and Pomeroy (2004). The functional form for $fSCA$ from Yang et al. (1997) could thus be inferred from integrating normal as well as log-normal snow depth distributions with subsequent fitting of the parametric

70   SCD curves. The main difference between the form of Yang et al. (1997) and Essery and Pomeroy (2004) is the variable in the denominator. Yang et al. (1997) used the aerodynamic roughness length whereas Essery and Pomeroy (2004) obtained the standard deviation of snow depth (σ_{HS}) at peak of winter in the denominator. The advantage of introducing σ_{HS} in the closed form for $fSCA$ is that subgrid terrain characteristics, contributing to shape the dominant spatial snow depth distribution, can be used to parameterize σ_{HS} and thus to extend the $fSCA$ parameterization of Essery and Pomeroy (2004) for mountainous

75   terrain (Helbig et al., 2015).

Until recently, it was not possible to derive an empirical parameterization for σ_{HS} based on high-resolution snow depth data due to the lack of such high-resolution spatial data. New measurement methods such as terrestrial laser scanning (TLS), airborne laser scanning (ALS) and airborne digital photogrammetry (ADP) nowadays provide a wealth of spatial snow data at fine-scale horizontal resolutions. Since recently, digital photogrammetry can also be applied to high-resolution optical satellite

80   imagery (Marti et al., 2016; Deschamps-Berger et al., 2020; Eberhard et al., 2020; Shaw et al., 2020). Snow depth data at these high resolutions now ~~allow to statistically analyze~~ enable statistical analyses of spatial snow depth patterns for various purposes (e.g. Melvold and Skaugen, 2013; Grünewald et al., 2013; Kirchner et al., 2014; Grünewald et al., 2014; Revuelto et al., 2014; Helbig et al., 2015; Voegeli et al., 2016; López-Moreno et al., 2017; Helbig and van Herwijnen, 2017; Skaugen and Melvold, 2019). Based on spatial snow depth data sets, σ_{HS} could be related to terrain parameters. For instance, Helbig et al.

85   (2015) parameterized σ_{HS} at peak of winter using spatial mean HS and subgrid terrain parameters, namely a squared slope related parameter and terrain correlation length, and Skaugen and Melvold (2019) parameterized σ_{HS} for the accumulation season using current spatial mean HS and stratifications according to landscape classes and standard deviations of squared slope. Though both approaches are promising and also somehow similar, e.g. both use the squared slope as significant scale variable, they also differ, e.g. in the considered horizontal scale lengths at the development of the parameterization. While the

90   parameterization of Helbig et al. (2015) was developed for squared grid cell sizes from 50 m to 3 km, Skaugen and Melvold (2019) presented parameterizations for 0.5 km x 1 km grid cells. Helbig et al. (2015) observed improved performances for larger

scales (> 1000 m), Skaugen and Melvold (2019) observed the same performances when validating it for 0.5 km x 10.25 km grid cells. This can be explained by the physical processes shaping the complex mountain snow cover predominantly interacting at different length scales with topography e.g. precipitation, wind and radiation (Liston, 2004). A multi-scale behaviour has been

95   found in various studies using different spatial coverages and measurement platforms (e.g. Deems et al., 2006; Trujillo et al., 2007; Schirmer et al., 2011; Mendoza et al., 2020), but a thorough analysis of spatial autocorrelations using many spatial snow depth data sets up to several kilometers in horizontal resolutions far below the first estimated scale break of about 10 to 20 m has not been presented so far. Such an analysis could reveal a scale range from which the spatial snow distribution in mountainous terrain can be parameterized with consistent accuracy. Using the newly available wealth of spatial snow data we now have the

100   opportunity to better understand the differences in previous empirically developed closed-form $fSCA$ parameterizations by adding variability in evaluation data sets, i.e. by using data from different geographic regions, as well as by taking into account the spatial scale in scaling parameters.

    This article presents a new $fSCA$ parameterization for mountainous terrain for various snow model applications. Since snow model applications operate at different spatial scales a $fSCA$ parameterization should work across spatial scales as well

105   as for various snow climates. Two important points were therefore tackled compared to a previous $fSCA$ parameterization: 1) We derived the empirical parameterization for σ_{HS} on a large pool of spatial snow depth data sets at peak of winter from various geographic sites and validated it scale- as well as region-dependent. 2) Based on a spatial scale analysis we introduced scale-dependent parameters in the peak of winter parameterization of Helbig et al. (2015) for σ_{HS} such that the new $fSCA$ parameterization can be reliably applied for grid cell sizes starting at 200 m up to 5 km. While a seasonal $fSCA$ model

110 algorithm can be built using parameterized σ_{HS} at peak of winter, we need additional information on the history of previous HS and SWE values to mimic the real seasonal $fSCA$ evolution. We will present a seasonal $fSCA$ algorithm separately.

2   Data

We compiled 11 spatial snow depth data sets from seven different geographic sites in mountainous regions of Switzerland, France and the US, i.e. from two continents (Figure 1). These data sets have horizontal grid cell resolutions Δx between 0.1

115   m and 3 m and cover areas from 0.14 km^2 to 280 km^2. In addition to that, the snow depth data sets were acquired by five different remote sensing methods, i.e. using different platforms. The diversity of the data sets can be seen in Figure 2 showing the pdfs for snow depth, elevation and the squared slope related parameter μ (Helbig et al., 2015) which is described in Section 3.1. All snow depth data was gathered at the local approximate point in time when snow accumulations had reached their annual maximum. Except for the two snow depth data sets shown in Figure 3, the data sets have been published before or

120 the geographic location is described elsewhere. In the following all snow depth data sets are listed, grouped according to their mountain range.

2.1   Eastern Swiss Alps

We used snow depth data sets acquired by three different platforms above four different alpine sites in the eastern Swiss Alps.

Figure 1. The map shows the approximate location of the eleven spatial snow depth data sets. The colors of the trays indicate the region, measurement platform or acquisition date as presented in Figure 2.
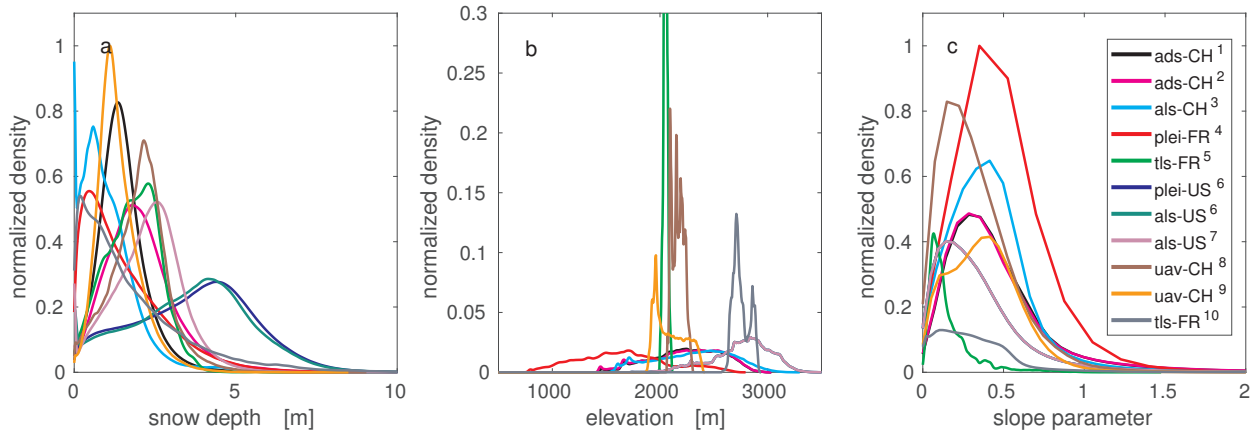
Figure 2. Probability density functions for fine-scale (a) snow depth, (b) elevation and (c) squared slope related parameter per observation data set in its original horizontal resolution, i.e. between 0.1 m and 3 m. Densities were normalized with the corresponding maximum density of all data sets. Note that for elevation (b) the y-axis was cut for better visibility. Colors represent the different geographic regions, measurement platform or acquisition dates (number) of the compiled data set as indicated in Section 2.1 to 2.4.

The first platform was airborne digital scanning (ADS) using an opto-electronic line scanner on an airplane. Data was acquired from the Wannengrat and Dischma area near Davos in the eastern Swiss Alps (Bühler et al., 2015). ADS-derived snow depth data sets were used from 20 March 2012 ('ads-CH^2') and 9 March 2016 ('ads-CH^1') together with summer digital elevation models (DEM) (Marty et al., 2019). The data set covers about 150 km^2 in 2 m resolution. Bühler et al. (2015) validated the 2 m ADS-derived snow depth data among others with TLS data. They obtained a root mean square error (RMSE) of 33 cm and a normalized median absolute deviation (NMAD) of the residuals (Höhle and Höhle, 2009) of 26 cm.

The second platform was an unmanned aerial system (UAS) recording optical imagery with real time kinematik (RTK) positioning of the image acquisition points of the snow cover by a standard camera over two different smaller regions near Davos in the eastern Swiss Alps (Bühler et al., 2016; Eberhard et al., 2020). These images were photogrammetrically processed into a digital surface model DSM. By subtracting the snow free DSM from the summer flight, the HS values were obtained (Bühler et al., 2017). An UAS-derived snow depth data set was used from 7 April 2018 ('uav-CH^9') from Schürlialp together with a UAS-acquired summer DEM (Eberhard et al., 2020). The Schürlialp data set covers about 3.2 km^2 which we used in 30 cm resolution. A second UAS-derived snow depth data set was used from 29 March 2019 ('uav-CH^8') from Gaudergrat together with a UAS-acquired summer DEM. The Gaudergrat data set covers about 0.8 km^2 in 10 cm resolution (Figure 3b). Compared to snow depth data from snow probing, Eberhard et al. (2020) obtained a RMSE of 16 cm and a NMAD of 11 cm for UAS-derived snow depth data at 9 cm horizontal resolution from Schürlialp.

The third platform was airborne laser scanning (ALS) above the Dischma region near Davos in the eastern Swiss Alps (Figure 3a). This acquisition was a Swiss partner mission of the Airborne Snow Observatory (ASO) (Painter et al., 2016). For consistency reasons, the same lidar setup was used and similar processing standards to the ASO campaigns in California were applied (Section 2.2). ALS-derived snow depth data was used from 20 March 2017 ('als-CH^3') together with a summer DEM from 2017. The ALS data set from Switzerland used here covers about 260 km^2 in 3 m resolution. Details on the derivation of the ALS data can be found in Mazzotti et al. (2019) though this study focused on three 0.5 km^2 forested sub data sets. Validation of 1 m ALS-derived snow depth grids from 20 March 2017 against data from snow probing within forest but outside canopy (i.e. not below a tree) resulted in a RMSE of 13 cm and a bias of -5 cm.

2.2 Sierra Nevada, CA, US

We used data sets acquired by two different platforms above Tuolumne basin in the Sierra Nevada (California) in the US.

The first platform was ALS performed by ASO (Painter et al., 2016). ALS-derived snow depth data was used from 26 March 2016 ('als-US^7') and 2 May 2017 ('als-US^6') together with a summer DEM (Painter, 2018). The second platform was a Pléiades product from 1 May 2017 ('plei-US^6'). A detailed data description of the Pléiades data set derivation is given in Deschamps-Berger et al. (2020).

We used the ASO summer DEM for the Pléiades as well as the ALS snow depth data sets. Given that the extent of the Pléiades snow depth data set was much smaller than the ALS domain, we cropped the ALS data sets to the Pléiades data set extension resulting in a coverage of about 280 km^2. The horizontal resolution used here was 3 m for both data sets. Compared to snow probe measurements in relatively flat areas ALS snow depth data at 3 m horizontal resolution was found unbiased with
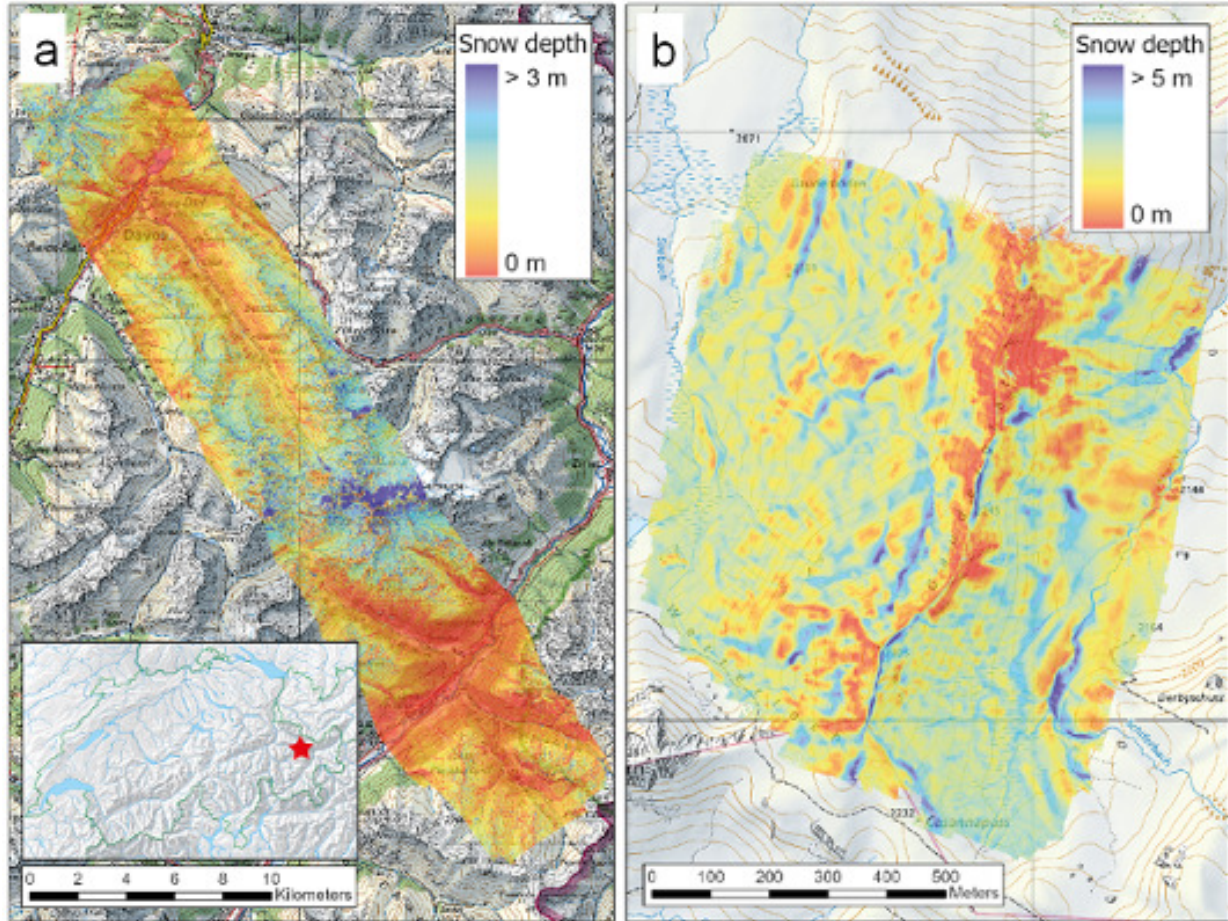
Figure 3. Snow depth maps of the eastern Swiss Alps: (a) in the Dischma region (ALS data) and (b) at Gaudergrat (UAS data) at peak of winter. The red dot in the inset map for Switzerland shows the location of the two sites. Pixmap © 2020 Swisstopo (5704000000), reproduced by permission of Swisstopo (JA100118).

a RMSE of 8 cm (Painter et al., 2016). Pléiades-derived snow depth data was recently validated with ASO data over 137 km^2 at 3 m resolution above Tuolumne basin (Deschamps-Berger et al., 2020). A RMSE of 80 cm, a NMAD of 69 cm and a mean bias of 8 cm was obtained for the Pléiades data set.

2.3 Eastern French Pyreenes

A Pléiades product was acquired over the Bassiès basin in the northeastern French Pyreenes. Pléiades-derived snow depth data was used from 15 March 2017 ('plei-FR^4') together with a summer DEM (Marti et al., 2016). The data set we used, covers about 113 km^2 in 3 m resolution. Marti et al. (2016) derived a median of the bias between 2 m Pléiades data and snow probe measurements of -16 cm and with UAS measurements of -14 cm. They further obtained a NMAD of 45 cm with snow probe measurements and a NMAD of 78 cm with UAS measurements.

2.4 Southeastern French Alps

TLS-derived snow depth data was acquired at two alpine mountain passes in the southeastern French Alps. One snow depth data set was acquired over Col du Lac Blanc at 9 March 2015 ('tls-FR^{10}') (Revuelto et al., 2020). A site and data description can be found in Naaim-Bouvet et al. (2010); Vionnet et al. (2014); Schön et al. (2015, 2018). We used a UAS-acquired summer DEM (Guyomarc'h et al., 2019). The data set covers about 0.6 km^2 in 1 m resolution. The second TLS-derived snow depth data set was acquired over Col du Lautaret at 27 March 2018 ('tls-FR^5') (Revuelto et al., 2020, under review). We used a TLS-acquired summer DEM. The data set covers about 0.14 km^2 in 1 m resolution. Previously, mean biases between 4 and 10 cm for TLS laser target distances up to 500 m were obtained between TLS-derived and reference tachymetry measurements (Prokop, 2008; Prokop et al., 2008; Grünewald et al., 2010).

2.5 Preprocessing

In all data sets grid cells Δx with forest, rivers, glaciers or buildings were masked out. In order to avoid introducing any biases we consistently neglected fine-scale snow depth values in all data sets that were lower than 0 m or larger than 15 m. We used a snow depth threshold of 0 m to decide whether or not a fine-scale grid cell was snow-covered.

3 Methods

We parameterize the standard deviation of snow depth σ_{HS} to reassess the validity of the $fSCA$ parameterization for complex topography of Helbig et al. (2015) for a range of spatial scales, in particular for sub-kilometer spatial scales.

3.1 Fractional snow-covered area parameterization (Helbig et al., 2015)

Helbig et al. (2015) derived a $fSCA$ parameterization by integrating a normal pdf assuming spatially homogeneous melt. Subsequent fitting over a range of coefficients of variation CV (standard deviation divided by its mean) between 0.06 and 1.00
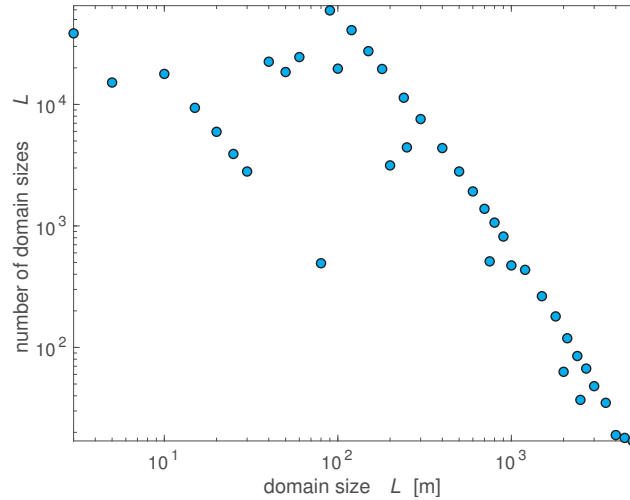
Figure 4. Total number of valid domain sizes L per domain size L in log-log scale.

resulted in a similar closed form fit for $fSCA$ as Essery and Pomeroy (2004) obtained by integrating a lognormal pdf:

$$fSCA = \tanh(1.3\frac{HS}{\sigma_{HS}}),\tag{1}$$

using current HS and standard deviation of previous maximum snow depth or peak of winter. The standard deviation of snow depth at peak of winter was derived by relating peak of winter high-resolution spatial snow depth data from Switzerland and Spain to underlying summer terrain parameters (Helbig et al., 2015)

$$\sigma_{HS} = HS^a \mu^b \exp[-(\xi/L)^2]\tag{2}$$

with $a = 0.549$, $b = 0.309$ and HS and terrain correlation length ξ in meters. ξ and μ are summer terrain parameters, where μ is related to the mean squared slope via $\mu = \left\{\overline{[(\partial_x z)^2 + (\partial_x z)^2]}/2\right\}^{1/2}$ using partial derivatives of subgrid terrain elevations z, i.e. from a DEM. The correlation length ξ or typical width of topographic features in a domain size L was derived via $\xi = \sqrt{2}\sigma_z/\mu$ with the standard deviation of elevations σ_z. The L/ξ ratio indicates how many characteristic topographic features of length scale ξ are included in each L. Figure 2 in Helbig et al. (2009) shows a transect of a topography indicating the described characteristic length scales. Similar to Helbig et al. (2015), we linearly detrended the summer DEM before deriving the terrain parameters to unveil the correct terrain characteristics associated with the shaping process of the snow depth distribution at the corresponding scale. Using Eq. (1), $fSCA$ can thus be derived with grid cell mean snow depth from a snow model and grid cell mean subgrid terrain parameters derived from a fine-scale summer DEM. With the σ_{HS} formulation shown in Eq. (2), Helbig et al. (2015) extended the $fSCA$ parameterization (Eq. (1)) for mountainous terrain.
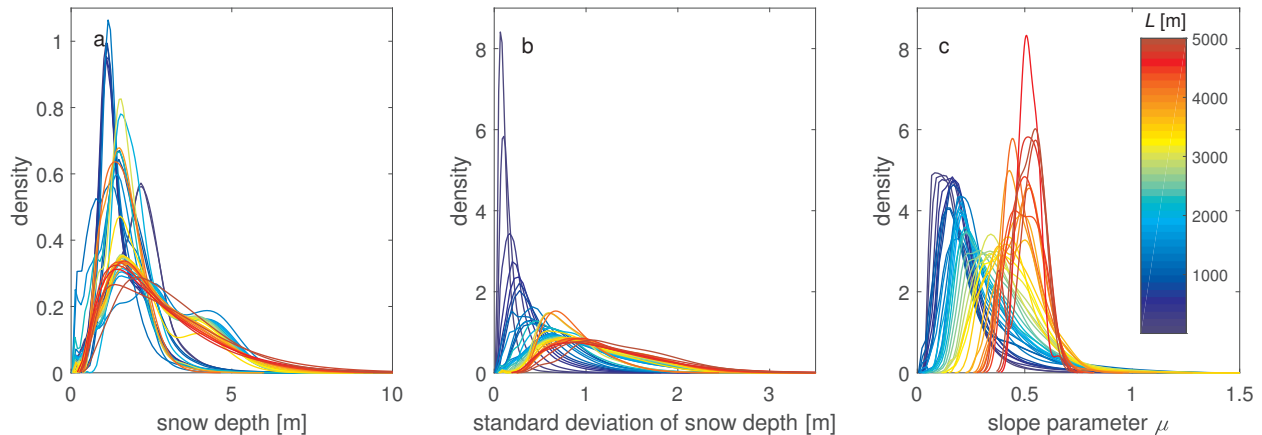
Figure 5. Probability density functions for (a) snow depth, (b) standard deviation of snow depth and (c) squared slope related parameter per domain size L after preprocessing, pooling all data sets and building spatial averages.

3.2 Aggregating and pooling of data sets

Pooling all snow depth data sets yields a data pool with a vast variety in snow climates, topographic characteristics and thus snow depth distributions. We first aggregated all snow data in squared ~~so-called~~ domain sizes L in regular grids between 3 m
205 to 5 km covering each geographic site. Our choice of the smallest applicable L in a data set was defined by a large enough $L/\Delta x$ ratio (here ≥ 20) to minimize the influence of grid cell resolutions when spatially averaging (Helbig et al., 2009). When aggregating, we required at least 70 % valid data in a domain size which was the maximum threshold to obtain a sufficient number of domains for the largest domain sizes L of 3 m to 5 km. In addition to that, we excluded L with spatial mean slope angles larger than 60 ° and spatial mean snow depth HS lower than 5 cm. By applying these limitations and since horizontal
210 resolutions Δx as well as the overall extent of the data sets vary, the full range of L, consisting of 41 different L values, was not represented by each data set. Overall, this resulted in a pool of 367'643 domains with L between 3 m and 5 km. We obtain a decreasing number of domains for increasing L with a range between 59'376 for $L = 90$ m and 17 for $L = 5000$ m (Figure 4). Spatial averages and standard deviations were built for each L. The resulting pooled data set shows a large variability in summer terrain characteristics. Spatial average slope angles range from 4° to 60° (μ from 0.05 to 1.22; Figure 5c), terrain
215 correlation lengths ξ from 6 m to 775 m and L/ξ-ratios from 3 to 40. Thus, typical summer terrain characteristics captured by coarse climate model grid cells are well represented. The diversity of the remaining domains with regards to snow depth is shown by means of the pdfs for spatial mean HS and σ_{HS} as a function of domain size L in Figure 5a and b. Since the data pool also covers a broad range in spatial mean HS (from 5 cm to 12.4 m) and spatial variability of snow depth σ_{HS} (from 1 cm to 4.6 m) (Figure 5a and 5b), we assume interannual snow depth variability is well described. In the following, overbars are
220 neglected for spatial averages, i.e. for instance HS represents spatial mean snow depth exclusively.
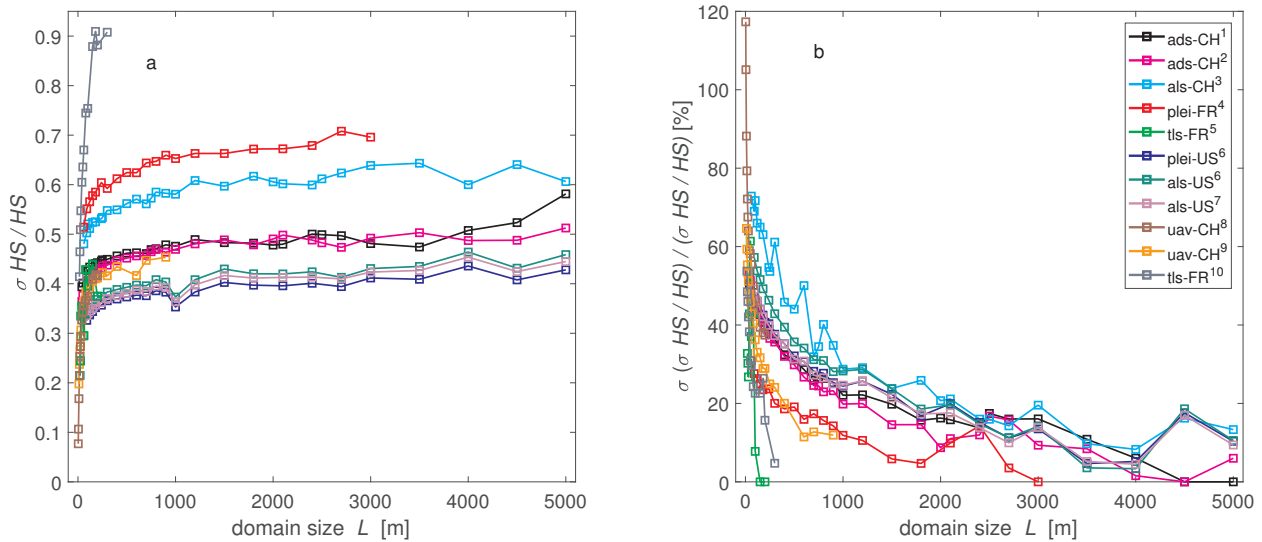
Figure 6. (a) Normalized standard deviation of HS as a function of domain size L for each data set separately. (b) Normalized percentage of standard deviation of (a) among each L. L ranges from 3 m to 5 km. Colors represent the different geographic regions, measurement platform or acquisition dates (number) of the compiled data set as indicated in Section 2.1 to 2.4.

3.3 Autocovariances for scale breaks

The spatial autocovariance allows finding spatial scale breaks up to which snow depth values are highly correlated, i.e. up to which length scale the snow depth distribution is strongly dictated by local topographic interactions of the snow cover with wind, precipitation and radiation. Below this scale break process models should ideally explicitly resolve these interactions to
225 reliably describe the spatial snow depth distribution. Above this scale break we assume that dominant wind or precipitation patterns due to larger scale topography impacts dictate spatial snow depth distributions. At this scale range the normalized standard deviations of snow depth σ_{HS} start levelling out (Figure 6a) as well as the normalized variability of σ_{HS} among similar sized L (Figure 6b).

We calculated spatial autocovariances for snow depth data sets with the "Fast Fourier Transform" (FFT), which allows
230 computing spatial autocovariances up to large distances by keeping the fine grid cell resolutions. We used the R function fft() of the 'stats' package (see R Core Team, 2020). For each autocovariance we then determine scale breaks using the R function uik() of the 'inflection' package (R Core Team, 2020).

3.4 ~~Fractional snow-covered area parameterization (Helbig et al., 2015)~~

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235 ~~Subsequent fitting over a range of coefficients of variation CV (standard deviation divided by its mean) between 0.06 and 1.00~~

resulted in a similar closed form fit for $fSCA$ as Essery and Pomeroy (2004) obtained by integrating a lognormal pdf:

$$fSCA = \tanh(1.3 \frac{HS}{\sigma_{HS}}),$$

using current HS and standard deviation of previous maximum snow depth or peak of winter. The standard deviation of snow depth at peak of winter was derived by relating peak of winter high-resolution spatial snow depth data from Switzerland and

240   Spain to underlying summer terrain parameters (Helbig et al., 2015)

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with $a = 0.549$, $b = 0.309$ and HS and terrain correlation length ξ in meters. ξ and μ are summer terrain parameters, where μ is related to the mean squared slope via $\mu = \left\{ \overline{[(\partial_x z)^2 + (\partial_y z)^2]}/2 \right\}^{1/2}$ using partial derivatives of subgrid terrain elevations z, i.e. from a DEM. The correlation length ξ or typical width of topographic features in a domain size L was derived

245   via $\xi = \sqrt{2}\sigma_z/\mu$ with the standard deviation of elevations σ_z. The L/ξ ratio indicates how many characteristic topographic features of length scale ξ are included in each L. Figure 2 in Helbig et al. (2009) shows a transect of a topography indicating the described characteristic length scales. Similar to Helbig et al. (2015), we linearly detrended the summer DEM before deriving the terrain parameters to unveil the correct terrain characteristics associated with the shaping process of the snow depth distribution at the corresponding scale. Using Eq. (1), $fSCA$ can thus be derived with grid cell mean snow depth from a

250   snow model and grid cell mean subgrid terrain parameters derived from a fine-scale summer DEM. With the σ_{HS} formulation shown in Eq. (2), Helbig et al. (2015) extended the $fSCA$ parameterization (Eq. (1)) for mountainous terrain.

3.4   Deriving a new scale-independent fractional snow-covered area parameterization

Helbig et al. (2015) showed that $fSCA$ performances increased with spatial scale and yielded best performance for spatial scales larger than 1000 m. Since the $fSCA$ parameterization was empirically developed on snow depth data from two geo-

255   graphic regions, here we reevaluated the scaling variables for the spatial variability of snow depth σ_{HS} as well as the functional form of the parameterization using the much larger compiled HS data set of this study. Various scaling variables were previously employed to capture σ_{HS} in mountainous terrain. Helbig et al. (2015) selected HS, the squared slope related parameter μ and the L/ξ ratio (Eq. (2)). Skaugen and Melvold (2019) used HS and standard deviation of the squared slope with sqS being derived using $sqS = tan^2 \zeta$, where ζ is the slope angle in radians. Since $tan^2 \zeta$ is the same than $2\mu^2$ (e.g. Löwe and Helbig,

260   2012), we here derive sqS from $2\mu^2$. Several other studies used σ_z as terrain parameter (e.g. Roesch et al., 2001). Here, we were interested in finding dominant scaling variables that correlate consistently across scales with σ_{HS}. We therefore analyzed the Pearson correlation coefficient r between various candidate parameters and σ_{HS} as a function of spatial scale, i.e. domain size L. Based on the results of previous studies, we selected the following candidate parameters: HS, μ, sqS, σ_{sqS}, L/ξ and σ_z.
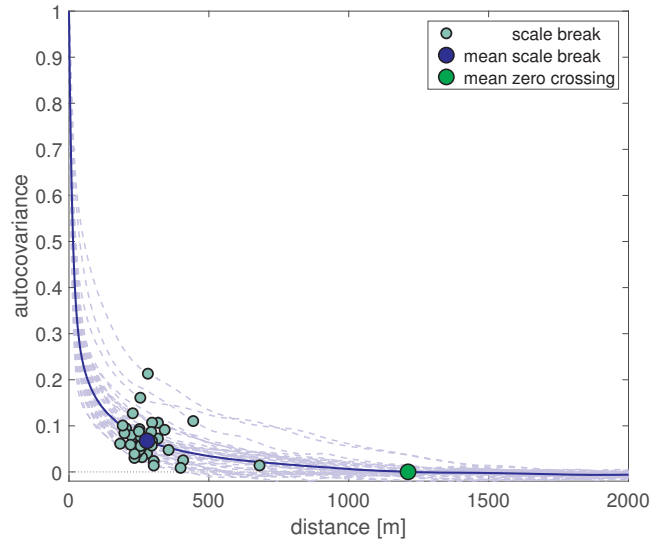
Figure 7. FFT derived autocovariances for spatial snow depth. Individual ranges, mean range and mean autocovariance zero crossing are shown.

3.5 Performance measures

The performance in this article is evaluated by the following measures: the root mean square error (RMSE), normalized root mean square error (NRMSE, normalized by the range of measured data (max-min) or the mean of the measurements for $fSCA$), mean absolute error (MAE), the mean absolute percentage error (MAPE, absolute bias with measured minus parameterized and normalized with measurements), the mean percentage error (MPE, bias with measured minus parameterized and normalized with measurements) and the Pearson correlation coefficient r as a measure for correlation. We also evaluate the performances by deriving the two-sample Kolmogorov-Smirnov test (K-S test) statistic values D (Yakir, 2013) for the pdfs and by computing the NRMSE for Quantile-Quantile plots (NRMSE$_{\text{quant}}$, normalized by the range of measured quantiles (max-min)) for probabilities with values in $[0.1, 0.9]$.

4 Results

4.1 Spatial correlation range from snow depth data

We derived a total of 40 autocovariances for the available domain sizes L of 3 km with grid cell sizes Δx of 2 or 3 m. We obtained scale breaks between 183 and 681 m with a mean of 284 m ($\pm\sigma$ 86 m) (Figure 7). The zero crossings for each autocovariance were between 402 m and 1815 m with a mean of 1011 m ($\pm\sigma$ 402 m). For the mean autocovariance we obtained a scale break at about 279 m and a zero crossing at about 1212 m. Based on the observed scale breaks we selected a minimum length scale of 200 m for deriving a new scale-dependent $fSCA$ parameterization for all larger scales. In the following all
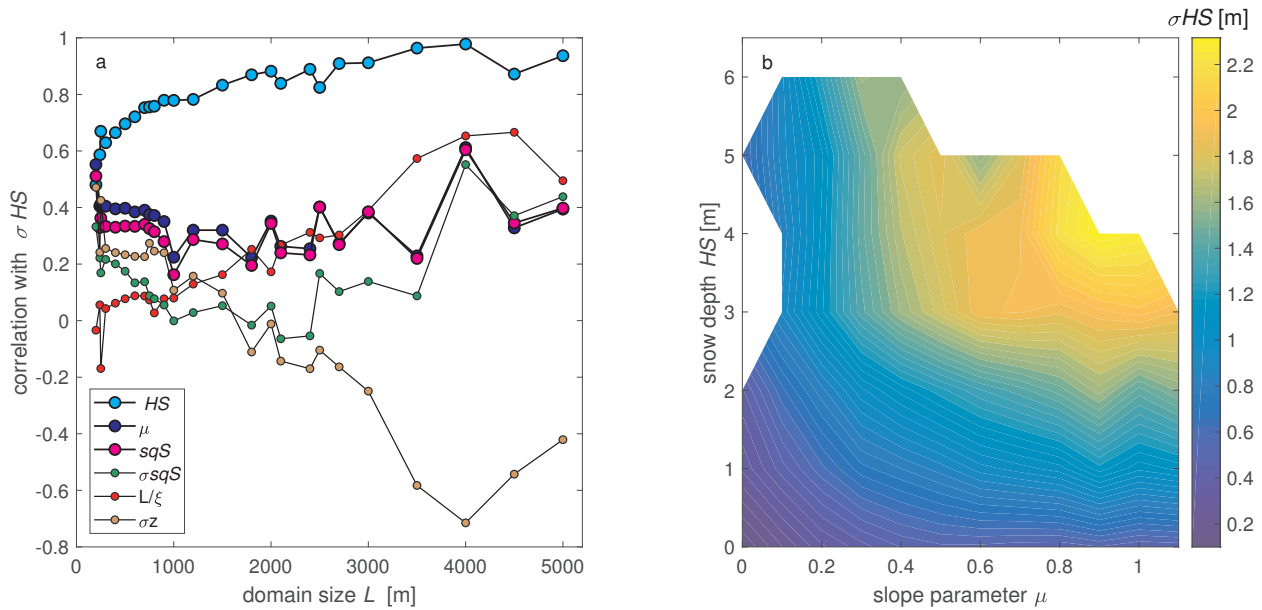
Figure 8. (a) Correlation coefficients between σ_{HS} and various parameters as a function of domain size L. (b) Standard deviation of snow depth σ_{HS} as a function of HS and μ.

results are therefore restricted to $L \geq 200$ m leaving a pool of 41'249 domain sizes L with L between 200 m and 5 km for the development of the parameterization.

4.2 Scaling variables for σ_{HS}

Correlation coefficients varied differently across spatial scales (Figure 8a). For all scales, we obtained the largest correlation
285 coefficients for HS ranging from 0.48 to 0.98 with a mean of 0.79. From correlations with the various subgrid terrain parameters, the largest correlations across all scales were reached for the squared slope related parameter μ ranging from 0.22 to 0.61 with a mean of 0.36. Similar consistent correlation coefficients across scales but slightly smaller for $L \leq 1800$ m resulted for the squared slope sqS with an overall mean of 0.33. The correlation coefficients for the standard deviation of sqS (σ_{sqS}) and σ_z were much less consistent across scales than for μ and sqS and were overall lower. The mean correlation for σ_{sqS} is
290 0.15, for L/ξ 0.21 and for σ_z 0.01. Though the mean correlation between σ_{HS} and L/ξ is rather low, correlation remains more consistent across scales up to about 2500 m and increases for larger scales considerably up to 0.67 (cf. Figure 8a).

We selected HS, μ and L/ξ as main scaling parameters for σ_{HS} across spatial scales from 200 m to 5 km (Figure 8b).

4.3 Scale-independent $fSCA$ parameterization

The correlation analysis across scales revealed the same dominant correlation parameters than in Helbig et al. (2015). We
295 therefore kept the functional form for σ_{HS} at peak of winter suggested by Helbig et al. (2015) using the three scaling variables
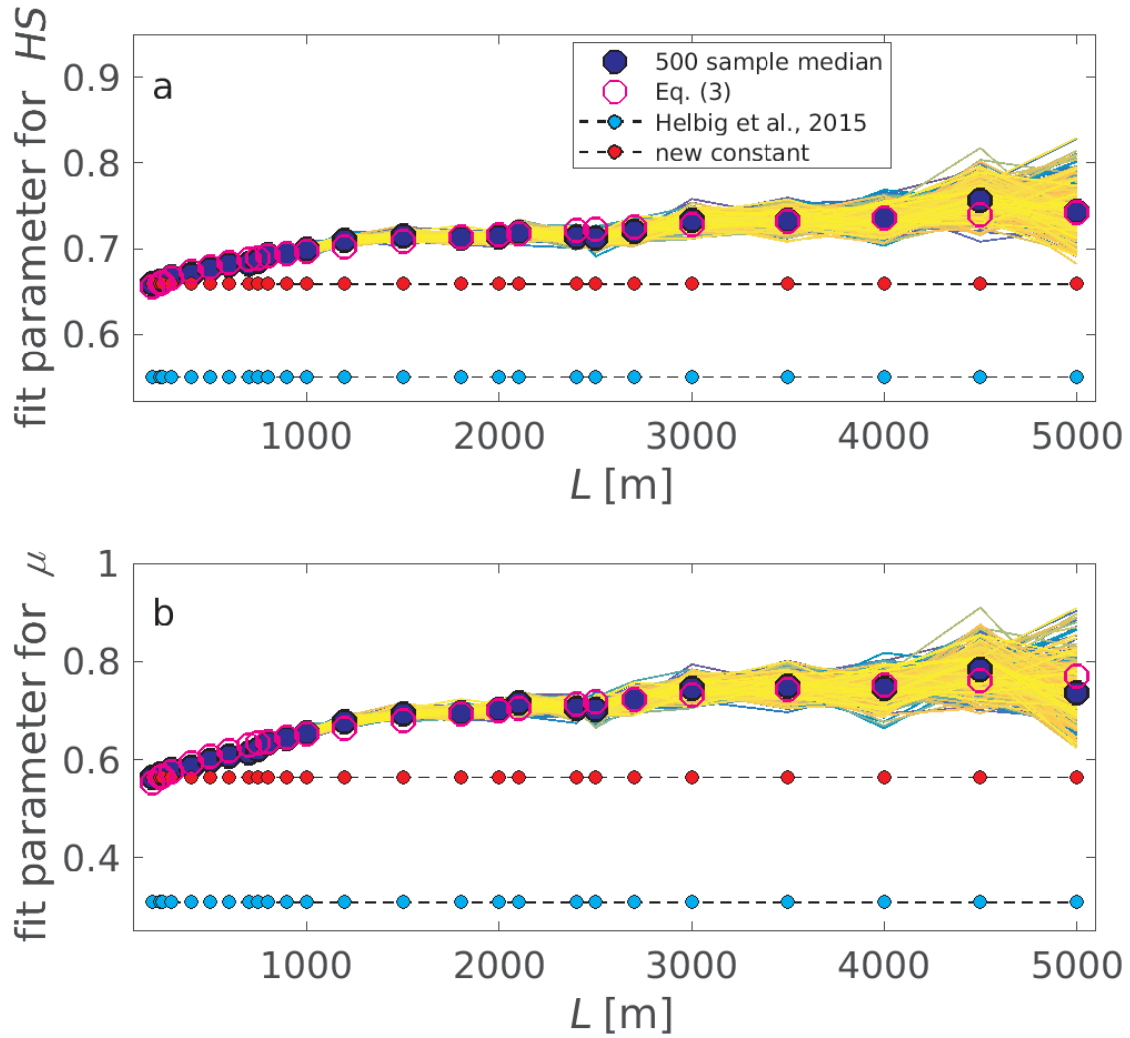
Figure 9. Fit parameters for Eq. (2) as a function of domain sizes L to scale variables (a) HS and (b) μ. Colored lines show the fit parameters derived for each of the 500 random 80 % samples of each of the 25 sub data pools. The dark blue dots depict the ensemble median per L. Previously obtained constant parameters of Helbig et al. (2015) (light blue dots) and newly fitted constant (red dots) as well as newly fitted scale-dependent (pink circles) parameters are shown.

HS, μ and L/ξ. The new σ_{HS} parameterization at peak of winter thus has the same functional form than the one suggested by Helbig et al. (2015) which was presented in Eq. (2). However, the fit parameters a and b therein are replaced by new parameters c and d which we specify below. To derive the new parameters c, d we fitted nonlinear regression models by robust M-estimators using iterated reweighed least squares (see R (R Core Team, 2020) and its robustbase package version 0.93-6 (Maechler et al., 2020)). We started at the scale length of 200 m, defined by the scale break which we derived before from spatial snow depth autocovariances.

Fit parameters were first derived for the entire data pool and $L \geq 200$ m yielding $c = 0.6589$ (± 0.0037) and $d = 0.5638$ (± 0.0043) with the 90 % confidence intervals of the fit parameters given in parentheses. These 'new' constant parameters c, d are larger than the previously derived constants a, b in Eq. (2) (cf. Figure 9).

In addition to fitting over the entire data pool and $L \geq 200$ m, we split the entire data pool into 25 sub data pools for any available domain size between 200 m to 5 km (cf. Figure 4). Thereby, each sub data pool comprised all domains larger or equal to the corresponding domain size, i.e. $L \geq 200$ m, $L \geq 240$ m etc. Fitting over such a sub data pool should allow describing the combined larger scale topography-wind-precipitation impacts on the spatial snow depth distribution in mountainous terrain acting at scales larger than the observed scale break of about 200 m. From each of the 25 created sub data pools, we randomly took 500 sub samples where each sub sample was restricted to 80 % data of the sub data pool. Each of the 500 sub samples per sub data pool was unique. Scale-dependent parameter values were derived for each of the 500 sub samples drawn from each of the 25 sub data pools (cf. individual colored lines in Figure 9). Given that the values of c, d clearly increase with spatial scale L (Figure 9) we introduced L in c, d to improve the application of Eq. (2) across scales. By fitting the ensemble medians of all scale-dependent fit parameters (dark blue dots in Figure 9) across all domain sizes between 200 m to 5 km, we obtained scale-dependent parameters $c(L)$ and $d(L)$. Thus, Eq. (2) using the following scale-dependent parameters $c(L)$ and $d(L)$ assembles our new σ_{HS} parameterization for $L \geq 200$ m:

$$c(L) = 0.5330\, L^{0.0389}$$
$$d(L) = 0.3193\, L^{0.1034}$$
(3)

with the 90 % confidence intervals of ± 0.0097, ± 0.0026 and ± 0.0183, ± 0.0079 in the order of introduced constants in Eq. (3).

The new σ_{HS} parameterization using $c(L)$ and $d(L)$ (Eq. (2) with Eq. (3)) is applied in the previously derived $fSCA$ parameterization (Eq. (1)). To demonstrate the resulting differences when using scale-dependent versus scale-independent fit parameters in parameterized σ_{HS} (Eq. (2) we will also validate the performance using constant c, d in the previously derived $fSCA$ parameterization as well as in the σ_{HS} parameterization.

4.4 Evaluation

4.4.1 Evaluation for σ_{HS} and $fSCA$ for all L

Parameterized σ_{HS} and $fSCA$ perform well for all domain sizes, i.e. for $L \geq 200$ m of the entire data pool. Very similar performance measures are obtained for the parameterizations using the newly derived constant fit parameters c, d and the

Table 1. Performance measures for all L between measurement and parameterization of (I) standard deviation of snow depth σ_{HS} with (a) Eq. (2) and constant or L dependent fit parameters c, d (Eq. (3)) and (b) σ_{HS} as in Helbig et al. (2015); Skaugen and Melvold (2019) and of (II) $fSCA$ with (a) Eq. (1) and (Ia) and (b) $fSCA$ as in Helbig et al. (2015) and Skaugen and Melvold (2019) using Eq. (1).

	NRMSE	RMSE	MPE	MAPE	MAE	r	K-S	NRMSE$_{quant}$
	[%]	[cm]	[%]	[%]	[cm]			[%]
I σ_{HS}								
(a) Eq. (2) with								
constant c, d parameter	7.9	26.6	-5.3	22.6	19.7	0.83	0.05	5.3
$c(L), d(L)$ (Eq. (3))	7.9	26.7	-4.1	22.4	19.6	0.83	0.05	5.5
(b) previous parameterizations from								
Helbig et al. (2015)	9.3	31.1	-29.5	36.7	25.3	0.82	0.22	14.6
Skaugen and Melvold (2019)	20.4	68.5	-77.9	82.8	57.9	0.68	0.48	37.6
	NRMSE	RMSE	MPE	MAPE	MAE	r	K-S	NRMSE$_{quant}$
	[%]		[%]	[%]				[%]
II $fSCA$								
(a) Eq. (1) with								
Eq. (2) and constant c, d parameter	2.4	0.02	0.22	1.11	0.01	0.64	0.37	0.5
Eq. (2) and $c(L), d(L)$ (Eq. (3))	2.4	0.02	0.16	1.09	0.01	0.63	0.37	0.4
(b) previous parameterizations from								
Helbig et al. (2015)	3.2	0.03	1.45	1.8	0.02	0.74	0.47	1.6
Skaugen and Melvold (2019) using Eq. (1)	6.2	0.06	3.87	4.8	0.05	-0.04	0.75	4.4

parameterizations using the scale-dependent parameters $c(L), d(L)$ (cf. Table 1 and I(a) and II(a)). We obtain a slightly better MPE for σ_{HS} when using scale-dependent fit parameters (-4 % versus -5 %) however for $fSCA$ MPEs are the same (0.2 %). The same rather low NRMSEs result for σ_{HS} (8 %) and for $fSCA$ (2 %).

4.4.2 Scale-dependent evaluation for σ_{HS} and $fSCA$

While mean performance measures of the σ_{HS} and $fSCA$ parameterization are almost uninfluenced to using constant or scale-dependent fit parameters (cf. Table 1 and I(a) and II(a)) we found diverging performances when analyzing performance measures as a function of scale (Figure 10). Across scales, improved or similar performances were achieved when using scale-dependent fit parameters in parameterized σ_{HS} especially for larger scales. Maximum performance improvements of 4 % occurred for L of 2500 m, respectively for $fSCA$ of 0.7 % when using scale-dependent fit parameters. Thus, introducing scale-dependent fit parameters enhanced the σ_{HS} parameterization for application across scales.
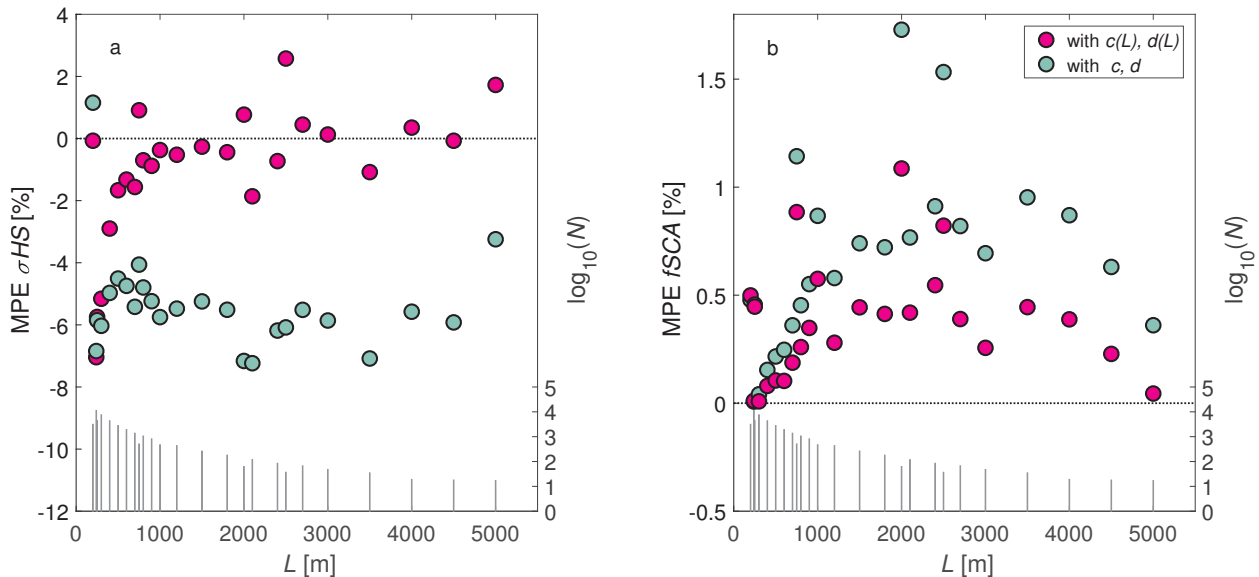
Figure 10. Mean percentage error (MPE) as a function of L for (a) σ_{HS} and (b) $fSCA$. MPEs are shown for the σ_{HS} and $fSCA$ parameterizations using Eq. (1) to (3) with scale-dependent $c(L), d(L)$ as well as for constant c, d. The second y-axis shows the number of valid domains per L on a logarithmic scale.

4.4.3 Scale- and region-dependent evaluation for σ_{HS} and $fSCA$

A large data set from various geographic regions allows us to develop a more reliable empirical parameterization than being
340 limited to the characteristics by a few data sets. Here, we not only compiled data sets from various geographic regions but the data sets were also acquired by different measurement platforms coming with a range in inaccuracies between below 10 cm to 80 cm. As a consequence larger scatter in performances appear when performance measures are depicted not only as a function of spatial scale but also region wise, including platform wise. While most of the MPEs are still between -20 % and 10 % some regions stand out because they have much larger MPEs when binned scale as well as region wise (Figure 11). For instance a
345 MPE of up to 60 % for σ_{HS} was obtained for TLS data from the southeastern French Alps and overall larger MPEs, though consistent across scales, for the Pléiades data from the northeastern French Pyreenes. MPEs for $fSCA$ on the other hand do not show a similar large spread among the regions and are low between -1 % to 2 % (Figure 11b).

4.4.4 Evaluation of previous closed form parameterizations

To increase our understanding of the performances achieved with the new parameterizations, we also tested two previously
350 derived empirical parameterizations. Specifically we investigated how parameterized σ_{HS} using Eq. (2) (Helbig et al., 2015) and using the recently published formulation of Skaugen and Melvold (2019) compare to observed σ_{HS} of our compiled data set (Figure 12a). We further tested both σ_{HS} parameterizations in the $fSCA$ parameterization (Eq. (1); Figure 12b). The parameterization of Helbig et al. (2015) works well. The performance measures for all L are only slightly worse compared
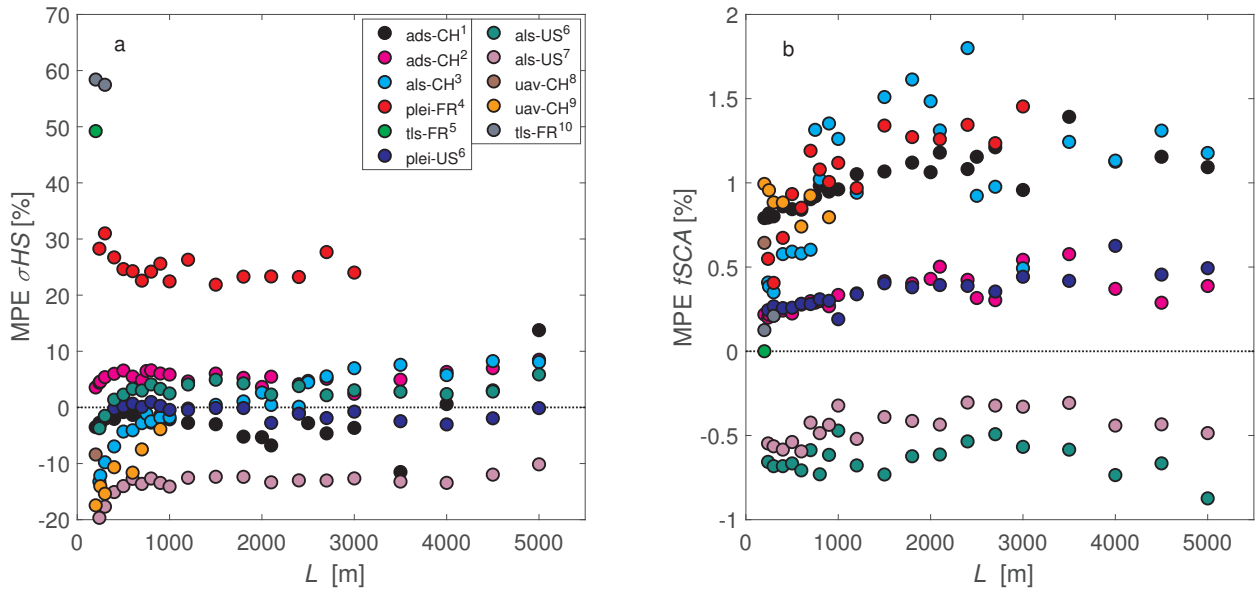
Figure 11. Mean percentage error (MPE) as a function of L for the compiled data set for (a) σ_{HS} and (b) $fSCA$ using Eq. (1) to (3) with scale-dependent $c(L), d(L)$. Colors represent the different geographic regions, measurement platform or acquisition dates (number) of the compiled data set as indicated in Section 2.1 to 2.4.

to the new parameterizations using both constant as well as scale-dependent fit parameters (Table 1). However, compared to the performance measures for the parameterization of Skaugen and Melvold (2019) the performances of Helbig et al. (2015) are clearly improved. Though MPEs of both previous σ_{HS} parameterizations are scale-dependent, the MPEs of Skaugen and Melvold (2019) reveal a larger scale-dependence of the performances compared to Helbig et al. (2015) (Figure 12a). In particular, individual MPEs vary a lot from MPEs for all L given in Table 1.

5 Discussion

5.1 Spatial correlation range

While multi-scale behaviour for spatial snow depth data has been found in various studies, observed scale breaks depend on the extent and horizontal resolution of the investigated snow depth data sets. A first scale break of spatial snow depth data in treeless, alpine terrain has been observed between 10 to 20 m (e.g. Deems et al., 2006; Trujillo et al., 2007; Schweizer et al., 2008; Schirmer and Lehning, 2011; Helfricht et al., 2014; Mendoza et al., 2020) and a second scale break has been observed at around 60 m (Trujillo et al., 2009). By computing spatial autocovariances starting with domain sizes L of 200 m in 0.1 m to 1 m resolution up to 3 km in 2 to 3 m resolution we also detected the two previously found scale breaks [not shown]. However, by additionally covering larger spatial extents than previously have been investigated, we also detected a third scale break with a mean at about 280 m (Figure 7). Similar long-range scale breaks between 185 m and 300 m were
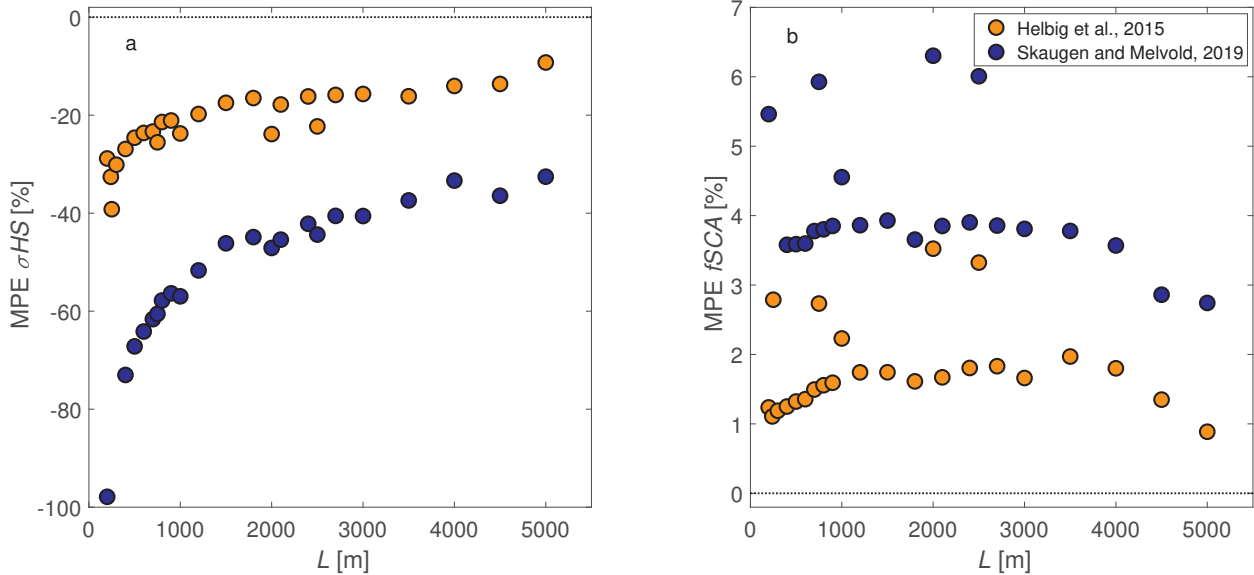
Figure 12. Mean percentage error (MPE) as a function of L for the compiled data set for (a) σ_{HS} and (b) $fSCA$. MPEs are shown for the σ_{HS} and $fSCA$ parameterizations of Helbig et al. (2015) as well as for the σ_{HS} parameterization of Skaugen and Melvold (2019) and Skaugen and Melvolds σ_{HS} parameterization applied in the $fSCA$ parameterization of Helbig et al. (2015) (Eq. (1)).

very recently reported from analyzing 24 TLS-derived snow depth data sets acquired during six snow seasons in a subalpine
370   catchment in the Spanish Pyrenees (Mendoza et al., 2020). Furthermore, a similar scale break at around 200 m was recently
found by analyzing performance decreases of distributed snow modelling in various grid cell sizes together with semivariogram
analysis of subgrid summer terrain slope angles in the same catchment in the High Atlas (Baba et al., 2019). While for other
application studies, such as in avalanche forecasting the smaller scale breaks are decisive for explicitly describing the relevant
snow cover processes, here we are more interested in the largest detected scale break. Above scale lengths of 200 m the longer
375   range processes of precipitation, wind and radiation interactions with topography most dominantly influence the spatial snow
distribution in mountainous terrain, while we believe there are different physical processes which establish the smaller scale
breaks at around 10 to 20 m and 60 m. The results presented here indicate that the model described by Eq. (1) and (3) is
reliably parameterizing the spatial snow distribution shaped by the longer range precipitation, wind and radiation interactions
with topography for spatial scales between 200 m and 5 km. Above the detected scale range of around 200 m not only the
380   spatial autocorrelations approach zero (Figure 7), but normalized σ_{HS} clearly start levelling out as well as the normalized
variability of σ_{HS} among similar sized L (Figure 6). Thus, even though we could not verify the $fSCA$ parameterization for
length scales larger than 5 km, we believe that as long as grid cell mean slope angles are larger than zero, Eq. (1) and (3) might
also hold for larger grid cell sizes than the 5 km.

5.2 Scaling parameter

385   We not only investigated dominant correlations between the spatial snow depth distribution and terrain parameters but we also analyzed these correlations as a function of spatial scale. For some commonly applied scaling parameters this revealed large variations of correlations across scales such as for σ_z (Figure 8a). Similar to our results, Skaugen and Melvold (2019) also obtained large correlations between σ_{HS} and mean squared slope sqS for spatial snow depth data sets acquired at peak of winter in Norway though this was only analyzed for grid cells of 0.5 km^2 (0.5 km x 1 km). Nevertheless, this confirms our

390   findings since mean squared slope is related to the slope related parameter μ used here by $sqS = 2\mu^2$. However, Skaugen and Melvold (2019) obtained slightly improved correlation for the standard deviation of squared slope and therefore selected this parameter to stratify the topography for parameterizing σ_{HS}. Across spatial scales as well as for all L we obtained lower correlations between the standard deviation of squared slope and σ_{HS} though we observed cross-correlations between mean and standard deviation of squared slope of 0.71 indicating that both parameters correlate well with σ_{HS}.

395 ## 5.3 Scale-independent $fSCA$ parameterization

The closed form fractional snow-covered area parameterization $fSCA$ given in Eq. (1) got enhanced by recalibration and introducing scale-dependent fit parameters (Eq. (3)) to make the performance consistent across spatial scales.

   We developed the parameterization on a large snow depth data set. Large variability in the snow depth data set was gained by compiling 11 individual data sets from varying geographic regions as well as various measurement platforms. While the

400   latter might explain remaining performance differences discussed below, the first led to large variability in summer terrain characteristics and snow climates and consequently spatial snow depth distributions (cf. Figure 2). Though our presented parameterization for σ_{HS} was empirically derived it is reassuring that for a new empirical derivation on a much larger and more diverse snow depth data set the same underlying functional form could be used. Furthermore, larger (about 17 % ~~respectively and~~ 45 %, respectively larger) but overall consistent constant fit parameters were obtained compared to those from Helbig et al.

405   (2015) based on a more limited number of data sets and just two geographic regions (cf. a, b in Eq. (2) and c, d presented in Section 4.3 or Figure 9).

   In addition to deriving constant fit parameters across spatial scales we took 500 random 80 % sub samples from each of the 25 sub data pools (Section 4.3). Scale-dependent constants considerably increased with increasing scale from L=200 m to L=5 km by at most 12 % ~~respectively~~ and 38 %, respectively (Figure 9). This demonstrates that accounting for scale-dependent

410   constants in the $fSCA$ parameterization (Eq. (1) with Eq. (2) and Eq. (3)) had to be performed. While we did not split our data set in development and validation subset, fitting over the ensemble median of all scale-dependent parameters to derive $c(L), d(L)$ ensures confidence in the resulting fit parameters.

   An increase in scatter among all $c(L)$ and $d(L)$ with increasing domain scale L (Figure 9) can be most likely explained by a concurrent decrease in available valid data in larger L. Though we required at least 70 % valid data per L when aggregating

415   fine-scale snow depth data in domain sizes L, the maximum threshold of 70 % was more often required for the larger L than for smaller L.

5.4 Evaluation

5.4.1 Evaluation for σ_{HS} and $fSCA$

Upon deriving performance measures on parameterized and observed σ_{HS} and $fSCA$ for all L (i.e. the pooled performance)
we obtained very similar performances when using newly derived constant or scale-dependent fit parameters, i.e. c, d or
$c(L), d(L)$ (Table 1). Despite considerable differences up to 12 % for c and up to 38 % for d between constant and scale-
dependent fit parameters (Figure 9) pooled performances for all L for σ_{HS} and $fSCA$ were similar (Table 1). An explanation
for this is that the number of available domains is strongly decreasing with increasing L. For $L \geq 3000$ m we have only about
0.33 % (137 in total) valid domains available compared to the total of 41'249 for $L < 3000$ m (Figure 4). This emphasizes the
need for a scale-dependent evaluation.

5.4.2 Scale-dependent evaluation for σ_{HS} and $fSCA$

The largest improvement in MPE for all L seem to originate from recalibration using the new compiled data set with a reduction
in MPE from -30 % to -5 % compared to a reduction from -5 % to -4 % when introducing scale-dependent fit parameters (Table
1). However, MPEs as a function of scale clearly demonstrated the improved behaviour when using scale-dependent $c(L), d(L)$
instead of constant fit parameters c, d in the σ_{HS} and $fSCA$ parameterization (Figure 10). Given that constant c, d were fitted
over the entire data set as have been $c(L), d(L)$, any performance improvement using $c(L), d(L)$ instead of constant c, d for
parameterized σ_{HS} and $fSCA$ originates in introducing scale-dependent parameters. For the parameterizations using the
constant fit parameters c, d errors varied slightly more across scales than when using the scale-dependent $c(L), d(L)$ version.
Individual scale-dependent errors were in part larger than the MPEs for all L given in Table 1. Unequal numbers of valid
domains per L most likely also contributed to this.

5.4.3 Scale- and region-dependent evaluation for σ_{HS} and $fSCA$

While we did not perform an evaluation using independent snow depth data sets, studying region-wise performances reveals the
spread in errors we can expect when the new parameterizations are applied on an individual independent data set (Figure 11).
We obtain much larger positive MPEs for σ_{HS} at lower spatial scales of $L = 200$ m and $L = 300$ m for the two TLS data sets
in the southeastern French Alps and overall larger MPEs between 20 to 30 %, though consistent across scales, for the Pléiades
data above the Bassiès basin in the northeastern French Pyrenees. It is unclear if these larger MPEs originate in uncertainties
of the data acquisition, i.e. are platform specific, or if they are linked to spatial snow depth distributions which could not be
captured by the proposed new parameterizations. RMSEs for the various remote sensing platforms and data sets used here
(Section 2) decrease from 80 cm for Pléiades data from the Sierra Nevada, to 33 cm for the ADS, to 16 cm for UAS, to 13
cm for ALS data from Switzerland, to 8 cm for ALS data from the Sierra Nevada and to 4 to 10 cm for TLS data in general.
Given the rather low errors typically obtained for TLS data compared to the other remote sensing platforms, the reason for the
large deviations of the TLS data sets might not originate in inaccuracies of the data acquisition. On the contrary, the observed

bias in the Pléiades data from the northeastern French Pyrenees might indeed be attributed to the rather large inaccuracies of the platform with NMADs of 45 cm to 78 cm (Marti et al., 2016). However, Pléiades data from the Sierra Nevada comes with a similar large NMAD of 69 cm but σ_{HS} can be parameterized very well with MPEs lower ±3 % across spatial scales (Figure 11a). Observed σ_{HS} from the TLS as well as from the Pléiades data in France was considerably larger than parameterized σ_{HS} but mean slope angles alone can also not explain this behaviour (between 6 and 23° for the TLS data and between 13 and 50° for the Pléiades data).

While we were not able to clearly relate some of the poorer region-wise performances to uncertainties related to the platform, other studies entirely focused on performing extensive inter-comparisons between platforms for large-scale snow depth mapping in alpine terrain (e.g. Bühler et al., 2015, 2016; Eberhard et al., 2020; Deschamps-Berger et al., 2020).

5.4.4 Evaluation of previous closed form parameterizations

Though we developed a new σ_{HS} parameterization (Eq. (3)), empirically derived parameterizations can only describe the variability inherent in the data set used to derive the parameterization. In addition to the region-wise evaluation, analyzing performances of previous empirically derived parameterizations may therefore allow estimating expected performance sensitivity to independent data sets. While both tested parameterizations of σ_{HS} (Helbig et al. (2015); Skaugen and Melvold (2019)) showed worse performances than the new parameterizations and less consistency as a function of scale, the model performances of Helbig et al. (2015) were only slightly worse than the new parameterizations (Table 1). The parameterization for σ_{HS} of Skaugen and Melvold (2019) was developed on mean domain sizes L of 750 m whereas Helbig et al. (2015) used Ls between 50 m to 3 km. This difference might be one reason for the overall poorer performances of Skaugen and Melvold (2019) compared to Helbig et al. (2015) across spatial scales. Since only one out of the 11 data sets used in this study was previously used to develop the parameterization of Helbig et al. (2015), an overall similar performance of Helbig et al. (2015) (Figure 12) with the large compiled data set of this study clearly confirms the underlying functional form of σ_{HS} suggested by Helbig et al. (2015) which was reapplied here.

6 Conclusions

We presented an empirical peak of winter parameterization for the standard deviation of snow depth σ_{HS} for treeless, mountainous terrain describing the spatial snow depth distribution in a grid cell for various model applications. The scaling variables of the new parameterization of σ_{HS} and $fSCA$ are the same than in Helbig et al. (2015) which are spatial mean snow depth, a squared slope related parameter and a terrain correlation length. All subgrid terrain parameters can be easily derived from fine-scale summer DEMs for each coarse grid cell.

By introducing spatial scale dependencies in the variables of the formulation for σ_{HS} of Helbig et al. (2015), σ_{HS} can be consistently parameterized across spatial scales starting at scales ≥200 m. The spatial snow depth variability or σ_{HS} is the important variable to parameterize the fractional-snow covered area $fSCA$ (Helbig et al., 2015). Performance improvements across spatial scales of the σ_{HS} parameterization therefore directly enhanced the $fSCA$ parameterization. Between length

scales of 200 m and 5 km mean percentage errors (MPE) were between -7 % and 3 % for σ_{HS} and between 0 % and 1 % for $fSCA$.

The subgrid parameterization of σ_{HS} was developed on 11 spatial snow depth data sets from 7 different geographic regions in high spatial resolutions between 0.1 m to 3 m and with spatial coverage between 0.14 to 280 km^2. An evaluation of two previously presented empirical σ_{HS} parameterizations confirmed the functional form of the parameterization of Helbig et al. (2015) as well as the need to enhance its performance across scales. By analyzing data from the large pool of spatial snow depth data sets, we were able to recalibrate the subgrid parameterization of σ_{HS} and achieved improved performances using new constant fit parameters. Additionally introducing a scale-dependency in the dominant scaling variables further improved the performance across spatial scales. Mean MPEs of σ_{HS} over all scales (i.e. pooled performance) reduced from -30 % using Helbig et al. (2015) to -5 % after recalibration to -4 % after introducing scale-dependent fit parameters (Table 1). Individual scale-dependent improvements in MPEs reached up to 4 % when using newly derived scale-dependent fit parameters compared to newly derived constant fit parameters for σ_{HS} on the large data pool. This shows the improvement thanks to introducing scale-dependent parameters (Figure 10). Towards estimating the possible spread in performances when applying empirically derived σ_{HS} and $fSCA$ for independent geographic regions we validated the parameterizations region- and scale-specific. While this clearly increased MPEs for three data sets, the majority of the region- and scale-dependent MPEs were between $\pm$ 10 % for σ_{HS} and between -1 % and 1.5 % for $fSCA$ indicating that the parameterizations perform similar well in most geographical regions.

By parameterizing the peak of winter σ_{HS} in mountainous terrain, Helbig et al. (2015) extended the $fSCA$ parameterization of Essery and Pomeroy (2004) for topography. Here, we extended the peak of winter σ_{HS} parameterization, and thus the $fSCA$ parameterization, to be applicable over a large range of snow climates and topographic characteristics as well as across spatial scales. Since $fSCA$ is an essential model parameter, a seasonal $fSCA$ algorithm describing the variability throughout a winter season is of relevance for various model applications. A peak of winter parameterization of σ_{HS} that describes the maximum spatial snow depth variability can however not alone be used to describe that seasonal $fSCA$ evolution. A reliable model application of any $fSCA$ parameterization requires an implementation accounting for alternating snow accumulation and melt events during the season, i.e. to describe the SCD. Especially at lower elevations the separation of the SCD in only one accumulation period followed by a melting period is no longer valid (Egli and Jonas, 2009). A description of an algorithm for a seasonal $fSCA$ model implementation which uses the new scale-independent peak of winter parameterization of σ_{HS} in the $fSCA$ parameterization presented here is currently in preparation. Extending the empirical peak of winter $fSCA$ parameterization to a broader range of spatial scales and snow climates was thus a meaningful step towards accounting for spatiotemporal variability in snow depth for multiple snow model applications.

Data availability. All data used in this study is described in the data section. The data can be downloaded from the referenced repositories or data availability is described in the referenced publications.

Author contributions. YB, LE, CDB, SG, MD, JR, JSD, TJ: data acquisition/processing, experiment/platform design and setup; NH: development of parameterization; All authors contributed to the paper with discussions and ideas. NH wrote the paper with contributions from all co-authors.

515 *Competing interests.* The authors declare that they have no conflict of interest.

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