

1 **Interannual snow accumulation variability on glaciers derived from repeat, spatially**
2 **extensive ground-penetrating radar surveys**

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11 **Abstract**

12 There is significant uncertainty regarding the spatiotemporal distribution of seasonal
13 snow on glaciers, despite being a fundamental component of glacier mass balance. To
14 address this knowledge gap, we collected repeat, spatially extensive high-frequency
15 ground-penetrating radar (GPR) observations on two glaciers in Alaska during the spring
16 of five consecutive years. GPR measurements showed steep snow water equivalent
17 (SWE) elevation gradients at both sites; continental Gulkana Glacier’s SWE gradient
18 averaged 115 mm 100 m⁻¹ and maritime Wolverine Glacier’s gradient averaged 440 mm
19 100 m⁻¹ (over >1000 m). We extrapolated GPR point observations across the glacier
20 surface using terrain parameters derived from digital elevation models as predictor
21 variables in two statistical models (stepwise multivariable linear regression and
22 regression trees). Elevation and proxies for wind redistribution had the greatest
23 explanatory power, and exhibited relatively time-constant coefficients over the study
24 period. Both statistical models yielded comparable estimates of glacier-wide average
25 SWE (1 % average difference at Gulkana, 4 % average difference at Wolverine),
26 although the spatial distributions produced by the models diverged in unsampled regions
27 of the glacier, particularly at Wolverine. In total, six different methods for estimating the
28 glacier-wide winter balance average agreed within ± 11 %. We assessed interannual
29 variability in the spatial pattern of snow accumulation predicted by the statistical models
30 using two quantitative metrics. Both glaciers exhibited a high degree of temporal
31 stability, with ~85 % of the glacier area experiencing less than 25 % normalized absolute
32 variability over this five-year interval. We found SWE at a sparse network (3 stakes per
33 glacier) of long-term glaciological stake sites to be highly correlated with the GPR-
34 derived glacier-wide average. We estimate that interannual variability in the spatial
35 pattern of winter SWE accumulation is only a small component (4–10 % of glacier-wide
36 average) of the total mass balance uncertainty and thus, our findings support the concept

37 that sparse stake networks effectively measure interannual variability in winter balance
38 on glaciers, rather than some temporally varying spatial pattern of snow accumulation.

39

40 **1. Introduction**

41 Our ability to quantify glacier mass balance is dependent on accurately resolving the
42 spatial and temporal distributions of snow accumulation and snow/ice ablation.

43 Significant advances in our knowledge of ablation processes have improved
44 observational and modelling capacities (Hock, 2005; Huss and Hock, 2015; Fitzpatrick et
45 al., 2017), yet comparable advances in our understanding of the distribution of snow
46 accumulation have not kept pace (Hock et al., 2017). Reasons for this discrepancy are
47 two-fold: (i) snow accumulation exhibits higher variability than ablation, both in
48 magnitude and length scale, largely due to wind redistribution in the complex high-relief
49 terrain where mountain glaciers are typically found (Kuhn et al., 1995) and (ii)
50 accumulation observations are typically less representative (i.e., one stake in a few
51 hundred meter elevation band) or less effective than comparable ablation observations
52 (i.e., precipitation gage measuring snowfall vs. radiometer measuring short-wave
53 radiation). This discrepancy presents a significant limitation to process-based
54 understanding of mass balance drivers. Furthermore, a warming climate has already
55 modified – and will continue to modify – the magnitude and spatial distribution of snow
56 on glaciers through a reduction in the fraction of precipitation falling as snow and an
57 increase in rain-on-snow events (McAfee et al., 2013; Klos et al., 2014; McGrath et al.,
58 2017; Beamer et al., 2017; Littell et al., 2018).

59

60 Significant research has been conducted on the spatial and, to a lesser degree, the
61 temporal variability of seasonal snow in mountainous and high-latitude landscapes (e.g.,
62 Balk and Elder, 2000; Molotch et al., 2005; Erickson et al., 2005; Deems et al., 2008;
63 Sturm and Wagner, 2010; Schirmer et al., 2011; Winstral and Marks, 2014; Anderson et
64 al., 2014; Painter et al., 2016). Although major advances have occurred in applying
65 physically-based snow distribution models (i.e., iSnobal (Marks et al., 1999), SnowModel
66 (Liston and Elder, 2006), Alpine 3D (Lehning et al., 2006)), the paucity of required
67 meteorological forcing data proximal to glaciers limits widespread application. Many
68 other studies have successfully developed statistical approaches that rely on the

69 relationship between the distribution of snow water equivalent (SWE) and physically-
70 based terrain parameters (also referred to as physiographic or topographic properties or
71 variables) to model the distribution of SWE across entire basins (e.g., Molotch et al.,
72 2005; Anderson et al., 2014; Sold et al., 2013; McGrath et al., 2015).

73

74 A major uncertainty identified by these studies is the degree to which these statistically
75 derived relationships remain stationary in time. Many studies (Erickson et al., 2005;
76 Deems et al., 2008; Sturm and Wagner, 2010; Schirmer et al., 2011; Winstral and Marks,
77 2014; Helfricht et al., 2014) have found ‘time-stability’ in the distribution of SWE,
78 including locations where wind redistribution is a major control on this distribution. For
79 instance, a climatological snow distribution pattern, produced from the mean of nine
80 standardized surveys, accurately predicted the observed snow depth in a subsequent
81 survey in a tundra basin in Alaska (~4–10 cm root mean square error (RMSE); Sturm and
82 Wagner, 2010). Repeat LiDAR surveys over two years at three hillslope-scale study plots
83 in the Swiss Alps found a high degree of correlation ($r=0.97$) in snow depth spatial
84 patterns (Schirmer et al., 2011). They found that the final snow depth distributions at the
85 end of the two winter seasons were more similar than the distributions of any two
86 individual storms during that two-year period (Schirmer et al., 2011). Lastly, an 11-year
87 study of extensive snow probing (~1200 point observations) at a 0.36 km² field site in
88 southwestern Idaho found consistent spatial patterns ($r=0.84$; Winstral and Marks, 2014).
89 Collectively, these studies suggest that in landscapes characterized by complex
90 topography and extensive wind redistribution of snow, spatial patterns are largely time-
91 stable or stationary, as long as the primary drivers are stationary.

92

93 Even fewer studies have explicitly examined the question of interannual variability in the
94 context of snow distribution on glaciers. Spatially-extensive snow probe datasets are
95 collected by numerous glacier monitoring programs (e.g., Bauder et al., 2017; Kjølmoen
96 et al., 2017; Escher-Vetter et al., 2009) in order to calculate a winter mass balance
97 estimate. Although extensive, such manual approaches are still limited by the number of
98 points that can be collected and uncertainties in correctly identifying the summer surface
99 in the accumulation zone, where seasonal snow is underlain by firn. One study of two

100 successive end-of-winter surveys of snow depth using probes on a glacier in Svalbard
101 found strong interannual variability in the spatial distribution of snow, and the
102 relationship between snow distribution and topographic features (Hodgkins et al., 2006).
103 Elevation was found to only explain 38–60 % of the variability in snow depth, and in one
104 year, snow depth was not dependent on elevation in the accumulation zone (Hodgkins et
105 al., 2006). Instead, aspect, reflecting relative exposure or shelter from prevailing winds,
106 was found to be a significant predictor of accumulation patterns. In contrast, repeat
107 airborne LiDAR surveys of a ~36 km² basin (~50% glacier cover) in Austria over five
108 winters found that the glacierized area exhibited less interannual variability (as measured
109 by the interannual standard deviation) than the non-glacierized sectors of the basin
110 (Helfricht et al., 2014). Similarly, a three-year study of snow distribution on
111 Findelgletscher in the Swiss Alps using ground-penetrating radar (GPR) found low
112 interannual variability, as 86 % of the glacier area experienced less than 25 % normalized
113 relative variability (Sold et al., 2016). These latter studies suggest that seasonal snow
114 distribution on glaciers likely exhibits ‘time-stability’ in its distribution, but few datasets
115 exist to robustly test this hypothesis.

116

117 The ‘time-stability’ of snow distribution on glaciers has particularly important
118 implications for long-term glacier mass balance programs, as seasonal and annual mass
119 balance solutions are derived from the integration of a limited number of point
120 observations (e.g., 3 to 50 stakes), and the assumption that stake and snow pit
121 observations accurately represent interannual variability in mass balance rather than
122 interannual variability in the spatial patterns of mass balance. Previous work has shown
123 ‘time-stability’ in the spatial pattern of annual mass balance (e.g., Vincent et al., 2017)
124 and while this is important for understanding the uncertainties in glacier-wide mass
125 balance estimates, the relative contributions of accumulation and ablation to this stability
126 are poorly constrained, thereby hindering a process-based understanding of these spatial
127 patterns. Furthermore, accurately quantifying the magnitude and spatial distribution of
128 winter snow accumulation on glaciers is a prerequisite for understanding the water budget
129 of glacierized basins, with direct implications for any potential use of this water, whether
130 that be ecological, agricultural, or human consumption (Kaser et al., 2010).

131

132 To better understand the ‘time-stability’ of the spatial pattern of snow accumulation on
133 glaciers, we present five consecutive years of extensive GPR observations for two
134 glaciers in Alaska. First, we use these GPR-derived SWE measurements to train two
135 different types of statistical models, which were subsequently used to spatially
136 extrapolate SWE across each glacier’s area. Second, we assess the temporal stability in
137 the resulting spatial distribution in SWE. Finally, we compare GPR-derived winter mass
138 balance estimates to traditional glaciological derived mass balance estimates and quantify
139 the uncertainty that interannual variability in spatial patterns in snow accumulation
140 introduces to these estimates.

141

142 **2. Study Area**

143 During the spring seasons of 2013 – 2017, we conducted GPR surveys on Wolverine and
144 Gulkana glaciers, located on the Kenai Peninsula and eastern Alaskan Range in Alaska
145 (Fig. 1). These glaciers have been studied as part of the U.S. Geological Survey’s
146 Benchmark Glacier project since 1966 (O’Neel et al., 2014). Both glaciers are ~16 km² in
147 area and span ~1200 m in elevation (426 – 1635 m asl for Wolverine, 1163 – 2430 m asl
148 for Gulkana). Wolverine Glacier exists in a maritime climate, characterized by warm air
149 temperatures (mean annual temperature = -0.2 °C at 990 meters; median equilibrium line
150 altitude for 2008 – 2017 is 1235 m asl) and high precipitation (median glacier-wide
151 winter balance = 2.0 m water equivalent (m w.e.)), while Gulkana is located in a
152 continental climate, characterized by colder air temperatures (mean annual temperature =
153 -2.8 °C at 1480 meters; median equilibrium line altitude for 2008 – 2017 is 1870 m asl)
154 and less precipitation (median glacier-wide winter balance = 1.2 m w.e.) (Fig. 2). The
155 cumulative mass balance time series for both glaciers is negative (~ -24 m w.e. between
156 1966–2016), with Gulkana showing a more monotonic decrease over the entire study
157 interval, while Wolverine exhibited near equilibrium balance between 1966 and 1987,
158 and sharply negative to present (O’Neel et al., 2014; O’Neel et al., 2018).

159

160 **3. Methods**

161 The primary SWE observations are derived from a GPR measurement of two-way travel
162 time (*tw*) through the annual snow accumulation layer. We describe five main steps to
163 convert *tw* along the survey profiles to annual distributed SWE products for each glacier.
164 These include (i) acquisition of GPR and ground-truth data, (ii) calculation of snow
165 density and associated radar velocity, which are used to convert measured *tw* to annual
166 layer depth and subsequently SWE, and (iii) application of terrain parameter statistical
167 models to extrapolate SWE across the glacier area. We then describe approaches to (iv)
168 evaluate the temporal consistency in spatial SWE patterns and (v) compare GPR-derived
169 SWE and direct (glaciological) winter mass balances.

170

171 **3.1. Radar data collection and processing**

172 Common-offset GPR surveys were conducted with a 500 MHz Sensors and Software
173 pulseEkko Pro system in late spring close to maximum end-of-winter SWE and prior to
174 the onset of extensive surface melt. GPR parameters were set to a waveform-sampling
175 rate of 0.1 ns, a 200-ns time window, and “Free Run” trace increments, where samples
176 are collected as fast as the processor allows, instead of at uniform temporal or spatial
177 increments.

178

179 In general, GPR surveys were conducted by mounting a plastic sled behind a snowmobile
180 and driving at a near-constant velocity of 15 km h⁻¹ (Fig. 3, S1, S2), resulting in a trace
181 spacing of ~20 cm. Coincident GPS data were collected using a Novatel Smart-V1 GPS
182 receiver (Omnistar corrected, L1 receiver with root-mean-square accuracy of 0.9 m
183 (Perez-Ruiz et al., 2011)). We collected a consistent survey track from year-to-year that
184 minimized safety hazards (crevasses, avalanche runouts) but optimized the sampling of
185 terrain parameter space on the glacier (e.g., range and distribution of elevation, slope,
186 aspect, curvature, etc.). However, in 2016 at Wolverine Glacier, weather conditions and
187 logistics did not allow for ground surveys to be completed. Instead, a number of radar
188 lines were collected via a helicopter survey. To best approximate the ground surveys
189 completed in other years, we selected a subset of helicopter GPR observations within 150
190 m of the ground-based surveys. Previous comparisons between ground and helicopter

191 platforms found excellent agreement in SWE point observations (coefficient of
192 determination (R^2)=0.96, root mean square error=0.14 m; McGrath et al., 2015).

193

194 Radargrams were processed using the ReflexW-2D software package (Sandmeier
195 Scientific Software). All radargrams were corrected to time zero, taken as the first
196 negative peak in the direct wave (Yelf and Yelf, 2006), and a dewow filter (mean
197 subtraction) was applied over 2 ns. When reflectors from the base of the seasonal snow
198 cover were insufficiently resolved, gain and band-pass filters were subsequently applied.
199 Layer picking was guided by ground-truth efforts and done semi-automatically using a
200 phase-following layer picker. For further details, please see McGrath et al. (2015).

201

202 **3.2. Ground truth observations**

203 We collected extensive ground-truth data to validate GPR surveys, including probing and
204 snowpit/cores. In the ablation zone of each glacier, we probed the snowpack thickness
205 every ~500 m along-track. In addition, we measured seasonal snow depth and density at
206 an average of five locations (corresponding to the glaciological observations; see Section
207 3.5) on each glacier in each year. Typically these locations include one or two in the
208 ablation zone, one near the long-term ELA, and two or more in the accumulation zone.
209 We measured snow density using a gravimetric approach in snowpits (at 10 cm intervals)
210 and with 7.25 cm diameter cores (if total depth >2 m; at 10–40 cm intervals depending on
211 natural breaks) to the previous summer surface. We calculated a density profile and
212 column-average density, ρ_{site} , at each site.

213

214 As snow densities did not exhibit a consistent spatial nor elevation dependency on the
215 glaciers (e.g., Fausto et al., 2018), we calculated a single average density, ρ , of all ρ_{site}
216 on each glacier and each year, which was subsequently used to calculate SWE:

217

$$218 \quad SWE = \left(\frac{twt}{2}\right) \cdot v_s \cdot \rho. \quad (1)$$

219

220 where twt is the two-way travel time as measured by the GPR and v_s is the radar
221 velocity. v_s was calculated for each glacier in each year as the average of two

222 independent approaches: (i) an empirical relationship based on the glacier-wide average ρ
223 (Kovacs et al., 1995) and (ii) a least-squares regression between snow depth derived by
224 probing and all radar *twt* observations within a 3-m radius of the probe site. An
225 exception was made at Wolverine in 2016 as no coincident probe depth observations
226 were made during the helicopter-based surveys. Instead, we estimated the second radar
227 velocity by averaging radar velocities calculated from observed *twt* and snow depths at
228 three snowpit/core locations.

229

230 **3.3. Spatial Extrapolation**

231 Extrapolating SWE from point measurements to the basin scale has been a topic of
232 focused research for decades (e.g., Woo and Marsh, 1978; Elder et al., 1995; Molotch et
233 al., 2005). Most commonly, the dependent variable SWE is related to a series of
234 explanatory terrain parameters, which are proxies for the physical processes that actually
235 control SWE distribution across the landscape. These include orographic gradient in
236 precipitation (elevation), wind redistribution of existing snow (slope, curvature, drift
237 potential), and aspect with respect to solar radiation and prevailing winds (eastness,
238 northness). We derived terrain parameters from 10-m resolution digital elevation models
239 (DEMs) sourced from the ArcticDEM project (Noh and Howat, 2015) for Gulkana and
240 produced from airborne Structure from Motion photogrammetry at Wolverine (Nolan et
241 al., 2015). Both DEMs were based on imagery from August 2015. Specifically, these
242 parameters include elevation, surface slope, surface curvature, northness (Molotch et al.,
243 2005), eastness, and snow drift potential (*Sb*) (Winstral et al., 2002; Winstral et al., 2013;
244 Fig. S3, S4). The *Sb* parameter is commonly used to identify locations where airflow
245 separation occurs based on both near and far-field topography and are thus likely
246 locations to accumulate snow drifts (Winstral et al., 2002). For specific details on this
247 calculation, please refer to Winstral et al. (2002). In the application of *Sb* here, we
248 determined the principle direction by calculating the modal daily wind direction during
249 the winter (October – May) when wind speeds exceeded 5 m s^{-1} (~minimum wind
250 velocity for snow transport; Li and Pomeroy, 1997). The length scales for curvature were
251 found using an optimization scheme that identified the highest model R^2 .

252

253 Prior to spatial extrapolation, we aggregated GPR observations to the resolution of the
254 DEM by calculating the median value of all observations within each 10 m pixel of the
255 DEM. We then utilized two approaches to extrapolate GPR point observations across the
256 glacier surface: (i) least-squares elevation gradient applied to glacier hypsometry and (ii)
257 statistical models. For (i), we derived SWE elevation gradients in two ways; first, solely
258 on observations that followed the glacier centerline and second, from the entire spatially-
259 extensive dataset. For (ii), we utilized two different models: stepwise multivariable linear
260 regressions and regression trees (Breiman et al., 1984). All of these approaches produced
261 a spatially-distributed SWE field over the entire glacier area. Individual points in this
262 field are equivalent to point winter balances (b_w ; m w.e.). From the distributed b_w field,
263 we calculated a mean area-averaged winter balance (B_w ; m w.e.).

264

265 Additionally, we implemented a cross-validation approach to the statistical models
266 (multivariable regression and regression tree), whereby 75 % of the aggregated
267 observations were used for training and 25 % were used for testing. However, rather than
268 randomly selecting pixels from across the entire dataset, we randomly selected a single
269 pixel containing aggregated GPR observations and then extended this selection out along
270 continuous survey lines until we reached 25 % of the total observational dataset, thus
271 removing entire sections (and respective terrain parameters) from the analysis (Fig. S5).
272 This approach provided a more realistic test for the statistical models, as the random
273 selection of individual cells did not significantly alter terrain-parameter distributions. For
274 each glacier and each year, we produced 100 training/test dataset combinations, but rather
275 than take the single model with the highest R^2 or lowest RMSE (between modelled SWE
276 and the GPR-derived test dataset), we produced a distributed SWE product by taking the
277 median value for each pixel from all 100 model runs and a glacier-wide median value
278 that is the median of all 100 individual B_w estimates. We chose the median-value
279 approach over a highest R^2 /lowest RMSE approach that is often utilized because, despite
280 being randomly selected, some training datasets were inherently advantaged by a more
281 complete sampling of terrain parameter distributions. These iterations resulted in the
282 highest R^2 /lowest RMSE when applied to the training dataset, but weren't necessarily

283 indicative of a better model, particularly in the context of being able to predict SWE at
284 locations on the glacier where the terrain parameter space had not been well sampled.

285

286 **3.3.2. Stepwise Multivariable Linear Regression**

287 We used a stepwise multivariable linear regression model of the form,

$$288 \text{SWE}_{(i,j)} = c_1x_{1(i,j)} + c_2x_{2(i,j)} + \dots + c_nx_{n(i,j)} + \varepsilon_{(i,j)}, \quad (2)$$

289 where $\text{SWE}_{(i,j)}$ is the predicted (standardized) value at location i,j and c_1, c_2, c_n are the beta
290 coefficients of the model, x_1, x_2, x_n are terrain parameters which are independent variables
291 that have been standardized and ε is the residual. We applied the regression model
292 stepwise and included an independent variable if it minimized the Akaike information
293 criterion (AIC; Akaike, 1974). We present the beta coefficients from each regression
294 (each year, each glacier) to explore the temporal stability of these terms.

295

296 **3.3.3. Regression Trees**

297 Regression trees (Breiman et al., 1984) provide an alternative statistical approach for
298 extrapolating point observations by recursively partitioning SWE into progressively more
299 homogenous subsets based on independent terrain parameter predictors (Molotch et al.,
300 2005; Meromy et al., 2013; Bair et al., 2018). The primary advantage of the regression
301 tree approach is that each terrain parameter is used multiple times to partition the
302 observations, thereby allowing for non-linear interactions between these terms. In
303 contrast, the MVR only allows for a single “global” linear relationship for each parameter
304 across the entire parameter-space. We implemented a random forest approach (Breiman,
305 2001) of repeated regression trees (100 learning cycles) in Matlab, using weak learners
306 and bootstrap aggregating (bagging; Breiman, 1996). Each weak learner omits 37% of
307 observations, such that these “out-of-bag” observations are used to calculate predictor
308 importance. The use of this ensemble/bagging approach reduces overfitting and thus
309 precludes having to subjectively prune the tree and provides more accurate and unbiased
310 error estimates (Breiman, 2001). Prior to implementing the regression tree, we removed
311 the SWE elevation gradient from the observations using a least-squares regression. As
312 described in the results, elevation is the dominant independent variable and as our
313 observations (particularly at Wolverine) did not cover the entire elevation range, the

314 regression tree approach was not well suited to predicting SWE at elevations outside of
315 the observational range.

316

317 **3.4. Interannual variability in spatial patterns**

318 We quantified the stability of spatial patterns in SWE across the five-year interval using
319 two approaches: (i) normalized range and (ii) the coefficient of determination. In the first
320 approach, we first divided each pixel in the distributed SWE fields by the glacier-wide
321 average, B_w , for each year and each glacier, and then calculated the range in these
322 normalized values over the entire five-year interval. For example, if a cell had normalized
323 values of 84 %, 92 %, 106 %, 112 % and 120 %, the normalized range would be 36 %. A
324 limitation of this approach is that it is highly sensitive to outliers, such that a single year
325 can substantially increase this range. This is similar to an approach presented by Sold et
326 al. (2016), but unlike their calculation (their Fig. 9), the normalized values reported here
327 have not been further normalized by the normalized mean of that pixel over the study
328 interval. Thus, the values reported here are an absolute normalized range, whereas Sold et
329 al. (2016) report a relative normalized range. In the coefficient of determination (R^2)
330 approach, we computed the least-squares regression correlation between the SWE in each
331 pixel and the glacier-wide average, B_w , derived from the MVR model over the five-year
332 period. For this approach, cells with a higher R^2 scale linearly with the glacier-wide
333 average, while those with low R^2 do not.

334

335 **3.5. Glaciological mass balance**

336 Beginning in 1966, glacier-wide seasonal (winter, B_w ; summer, B_s) and annual balances (B_a)
337 were derived from glaciological measurements made at three fixed locations on each glacier.
338 The integration of these point measurements was accomplished using a site-index method –
339 equivalent to an area-weighted average (March and Trabant, 1996; van Beusekom et al., 2010).
340 Beginning in 2009, a more extensive stake network of seven to nine stakes was established on
341 each glacier, thereby facilitating the use of a balance profile method for spatial extrapolation
342 (Cogley et al., 2011). Systematic bias in the glaciological mass balance time-series is removed
343 via a geodetic adjustment derived from DEM differencing over decadal timescales (e.g.,
344 O’Neel et al., 2014). For this study, glaciological measurements were made within a day of the

345 GPR surveys, and integrated over the glacier hypsometry using both the historically applied
346 site-index method (based on the long-term three stake network) and the more commonly
347 applied balance profile method (based on the more extensive stake network). We utilized a
348 single glacier hypsometry, derived from the 2015 DEMs, for each glacier over the entire five-
349 year interval. Importantly, in order to facilitate a more direct comparison to the GPR-derived
350 B_w estimates, we used glaciological B_w estimates that have not been geodetically calibrated.

351

352 **4. Results**

353 **4.1. General accumulation conditions**

354 Since 1966, Wolverine Glacier's median B_w (determined from the stake network) exceeds
355 Gulkana's by more than a factor of two (2.3 vs. 1.1 m w.e.), and exhibits greater
356 variability, with an interquartile range more than twice as large (0.95 m w.e. vs. 0.4 m
357 w.e.). Over the five-year study period, both glaciers experienced accumulation conditions
358 that spanned their historical ranges, with one year in the upper quartile (including the 5th
359 greatest B_w at Wolverine in 2016), one year within 25% of the median, and multiple years
360 in the lower quartile (2017 at Gulkana and 2014 at Wolverine had particularly low B_w
361 values) (Fig. 2). In all years, B_w at Wolverine was greater, although in 2013 and 2014, the
362 difference was only 0.1 m w.e.

363

364 Average accumulation season (taken as October 1 – May 31) wind speeds over the study
365 period were stronger (~ 7 m s^{-1} vs. ~ 3 m s^{-1}) and from a more consistent direction at
366 Wolverine than Gulkana (northeast at Wolverine, southwest to northeast at Gulkana)
367 (Fig. S6). On average, Wolverine experienced ~ 50 days with wind gusts > 15 m s^{-1} each
368 winter, while for Gulkana, this only occurred on ~ 7 days. Over the five-year study period,
369 interannual variability in wind direction was very low at Wolverine (2016 saw slightly
370 greater variability, with an increase in easterly winds). In contrast, at Gulkana, winds
371 were primarily from the northeast to east in 2013–2015, from the southwest to south in
372 2016–2017, and experienced much greater variability during any single winter.

373

374 **4.2. *In situ* and GPR point observations**

375 Glacier-averaged snow densities across all years were 440 kg m^{-3} (range $414\text{--}456 \text{ kg m}^{-3}$)
376 3) at Wolverine and 362 kg m^{-3} (range $328\text{--}380 \text{ kg m}^{-3}$) at Gulkana (Table S1). Average
377 radar velocities were 0.218 m ns^{-1} (range $0.207\text{--}0.229 \text{ m ns}^{-1}$) at Wolverine and 0.223 m
378 ns^{-1} ($0.211\text{--}0.231 \text{ m ns}^{-1}$) at Gulkana. Over this five-year interval, the GPR point
379 observations revealed a general pattern of increasing SWE with elevation, along with
380 fine-scale variability due to wind redistribution (e.g., upper elevations of Wolverine) and
381 localized avalanche input (e.g., lower west branch of Gulkana) (Fig. S1, S2). The
382 accumulation season (hereafter, winter) SWE elevation gradient was steeper (~ 440 vs.
383 $\sim 115 \text{ mm } 100 \text{ m}^{-1}$) and more variable in its magnitude at Wolverine than Gulkana.
384 Gradients ranged between $348\text{--}624 \text{ mm } 100 \text{ m}^{-1}$ at Wolverine, and $74\text{--}154 \text{ mm } 100 \text{ m}^{-1}$
385 1 at Gulkana (Fig. 4). Over all five years at both glaciers, elevation explained between 50
386 % and 83 % of the observed variability in SWE (Fig. 4).

387

388 **4.3. Model performance**

389 To evaluate model performance in unsampled locations of the glacier, both extrapolation
390 approaches were run 100 times for each glacier and each year, each time with a unique,
391 randomly selected training (75 % of aggregated observations) and test (remaining 25 %
392 of aggregated observations) dataset. The median and standard deviation of the
393 coefficients of determination (R^2) between modeled SWE and the test datasets for the 100
394 models runs are shown in Fig. 5. Model performance ranged from 0.25 to 0.75, but on
395 average, across both glaciers and all years, was 0.56 for the MVR approach and 0.46 for
396 the regression tree. Model performance was higher and more consistent at Wolverine,
397 whereas 2015 and 2017 at Gulkana had test dataset R^2 of ~ 0.4 and 0.3 , likely reflecting
398 the lower winter SWE elevation gradients and coefficients of determination with
399 elevation during these years (Fig. 4). The wide range in R^2 across the 100 model runs
400 reflects the variability in training and test datasets that were randomly selected. When the
401 test dataset terrain parameter space was captured by the training dataset, a high
402 coefficient of determination resulted, but when the test dataset terrain parameter space
403 was exclusive (e.g., contained only a small elevation range), the model performance was
404 typically low. This further highlights the importance of elevation as a predictor for these
405 glaciers.

406

407 At Gulkana, the model residuals (Fig. S1) exhibited spatiotemporal consistency, with
408 positive residuals (i.e., observed SWE exceeded modeled SWE by ~ 0.2 m w.e.) at mid-
409 elevations of the west branch, and at the very terminus of the glacier. The largest negative
410 residuals typically occurred at the highest elevations. In both cases, these locations
411 deviated from the overall SWE elevation gradient. At Wolverine, observations at the
412 highest elevations typically exceeded the modeled SWE (i.e., positive residuals),
413 particularly at the highest elevations of the northeast corner where wind drifting is
414 particularly prevalent (Fig. S2). For example, in 2015, nearly 80% of the residuals in this
415 section were positive and had a median value of 0.4 m. Elsewhere at Wolverine, the
416 residuals often alternated between positive and negative values over length scales of 10s
417 to 100s of meters (Fig. S2), which we interpret as zones of scour/drift not captured by the
418 MVR model.

419

420 The beta coefficients of terrain parameters from the MVR were fairly consistent from
421 year-to-year at both glaciers (Fig. 6). At Wolverine, elevation was the largest beta
422 coefficient, followed by Sb and curvature. At Gulkana, elevation was also the largest beta
423 coefficient, followed by curvature. Gulkana experiences much greater variability in wind
424 direction during the winter months (Fig. S6), possibly explaining why Sb was either not
425 included or had a very low beta coefficient in the median regression model. As our
426 surveys were completed prior to the onset of ablation, terrain parameters related to solar
427 radiation gain (notably the terms that include aspect: northness and eastness) had small
428 and variable beta coefficients.

429

430 **4.4. Spatial Variability**

431 A common approach for quantifying snow accumulation variability across a range of
432 means is the coefficient of variation (CoV), which is calculated as the ratio of the
433 standard deviation to the mean (Liston et al., 2004; Winstral and Marks, 2014). The mean
434 and standard deviation of CoVs at Wolverine were 0.42 ± 0.03 and at Gulkana, $0.29 \pm$
435 0.05 , indicating relatively lower spatial variability in SWE at Gulkana (Fig. 7). CoVs
436 were fairly consistent across all five years, although 2017 saw the largest CoVs at both

437 glaciers. Interestingly, 2017 had the lowest absolute spatial variability (i.e., lowest
438 standard deviation), but also the lowest glacier-wide averages during the study period,
439 resulting in greater CoVs.

440

441 Qualitatively, both Wolverine and Gulkana glaciers exhibited consistent spatiotemporal
442 patterns in accumulation across the glacier surface, with elevation exerting a first-order
443 control (Fig. 8, S7, S8). Overlaid on the strong elevational gradient are consistent
444 locations of wind scour and deposition, reflecting the interaction of wind redistribution
445 and complex – albeit relatively stable year to year – surface topography (consisting of
446 both land and ice topography). For instance, numerous large drifts (~2 m amplitude, ~200
447 m wavelength) occupy the northeast and northwest corners of Wolverine Glacier, where
448 prevailing northeasterly winds consistently redistributed snow into sheltered locations in
449 each year of the study period (Fig. 8). The different statistical extrapolation approaches
450 produced nearly identical B_w estimates (4 % difference on average at Wolverine and 1 %
451 difference on average at Gulkana) (Fig. 9). The MVR B_w estimate was larger in 4 out of 5
452 years at Wolverine (Fig. 9), while neither approach exhibited a consistent bias at
453 Gulkana.

454

455 Although the glacier-wide averages between these approaches showed close agreement,
456 we explored the differences in spatial patterns by calculating a mean SWE difference
457 map for each glacier by differencing the five-year mean SWE produced by the regression
458 tree model from the same produced by the MVR model (Fig. 10). As such, locations
459 where the MVR exceeded the regression tree are positive (yellow). At Gulkana, where
460 the two approaches showed slightly better glacier-wide B_w agreement, the magnitude in
461 individual pixel differences were substantially less than at Wolverine (e.g., color bar
462 scales range ± 0.2 m at Gulkana vs. ± 0.5 m at Wolverine). At Wolverine Glacier, there
463 were three distinct elevation bands where the MVR approach predicted greater SWE,
464 namely the main icefall in the ablation zone, a region of complex topography centered
465 around a normalized elevation of 0.65, and lastly, at higher elevations, where both
466 approaches predicted a series of drift and scour zones, although in sum, the MVR model
467 predicted greater SWE.

468

469 We used two different approaches to quantify the ‘time-stability’ of spatial patterns
470 across these glaciers. By the first metric, normalized range, we found that both glaciers
471 exhibited very similar patterns (Fig. 11), with either ~65 or 85 % (regression tree and
472 MVR, respectively) of the glacier area experiencing less than 25 % absolute normalized
473 variability (Fig. 12). The R^2 approach provides an alternative way of assessing the time
474 stability of SWE, essentially determining whether SWE at each location scales with the
475 glacier-wide value. By this metric, 80 % of the glacier area at Wolverine and 96 % of the
476 glacier area at Gulkana (based on MVR model) had a coefficient of determination greater
477 than 0.8 (Fig. 12), suggesting that most locations on the glacier have a consistent
478 relationship with the mean glacier-wide mass balance. By both metrics, the MVR output
479 suggests greater ‘time-stability’ (e.g., lower normalized range or higher R^2) compared to
480 the regression tree.

481

482 **4.5. Winter mass balance**

483 In order to examine systematic variations between the approaches we outlined in Section
484 3 for calculating the glacier-wide winter balance, B_w , we first calculated a yearly mean
485 from the six approaches (including four based on the GPR observations: MVR,
486 regression tree, elevation gradient derived from centerline only observations, elevation
487 gradient derived from all point observations, and two based on the *in situ* stake network:
488 site-index and profile). In general, Gulkana exhibited greater agreement (4 % average
489 difference) among the approaches, with most approaches agreeing within 5 % of the six-
490 approach mean (Fig. 13; Table S2). Wolverine showed slightly less agreement (7 %
491 average difference), as the two terrain parameters statistical extrapolations (MVR and
492 regression tree) produced B_w estimates ~9 % above the mean, while the two stake derived
493 estimates were ~7 % less than the mean. On average across all five years at Wolverine,
494 the MVR approach was the most positive, while the glaciological site-index approach
495 was always the most negative (Fig. 13). At both glaciers, the estimates using elevation as
496 the only predictor yielded B_w estimates on average within 3 % of the six-method mean,
497 with the centerline only based estimate being slightly negatively biased, and the complete
498 observations being slightly positively biased.

499

500 To examine the systematic difference between the glaciological site-index method and
501 GPR-based MVR approach, we compared stake-derived b_w values from the three long-
502 term stakes to all GPR-based MVR b_w values within that index zone (Fig. 14). Both the
503 stakes and the GPR-derived b_w values have been normalized by the glacier-wide value to
504 make these results comparable across years and glaciers. It is apparent that Wolverine
505 experienced much greater spatial variability in accumulation, with larger interquartile
506 ranges and a large number of positive outliers in all index zones. Importantly, the stake
507 weight in the site-index solution is dependent on the hypsometry of the glacier, and for
508 both glaciers, the upper stake accounts for ~65 % of the weighted average. In years that
509 the misfit between GPR B_w and site-index B_w was largest (2015 and 2016 at Gulkana,
510 2013 and 2017 at Wolverine), the stake-derived b_w at the upper stake was in the lower
511 quartile of all GPR-derived b_w values, explaining the significant difference in B_w
512 estimates in these years. Potential reasons for this discrepancy are discussed in Section
513 5.3.

514

515 *In situ* stake and pit observations traditionally serve as the primary tool for deriving
516 glaciological mass balances. However, in order for these observations to provide a
517 systematic and meaningful long-term record, they need to record interannual variability
518 in mass balance rather than interannual variability in spatial patterns of mass balance. To
519 assess the performance of the long-term stake sites, we examined the interannual
520 variability metrics for the stake locations. By both metrics (normalized absolute range
521 and R^2), the middle and upper elevation stakes at both glaciers appear to be in locations
522 that achieve this temporal stability, having exhibited ~10 % range and $R^2 > 0.95$ over the
523 five-year interval. The lower elevation stake was less temporally stable and exhibited
524 opposing behavior at each glacier. At Gulkana, this stake had a high R^2 (0.93) and
525 moderate normalized variability (26 %), which in part, reflects the lower total
526 accumulation at this site and the ability for a single uncharacteristic storm to alter this
527 total amount significantly. In contrast, Wolverine's lowest site exhibited both low R^2
528 (< 0.01) and normalized range (2 %), a somewhat unlikely combination. The statistical
529 models commonly predicted zero or near-zero cumulative winter accumulation at this site

530 (i.e., mid-winter rain and/or ablation is common at this site), so although the normalized
531 range was quite low, predicted SWE values were uncorrelated with B_w over the study
532 interval.

533

534 **Discussion**

535 **5.1. Interannual variability in spatial patterns**

536 Each glacier exhibited consistent normalized SWE spatial patterns across the five-year
537 study, reflecting the strong control of elevation and regular patterns in wind redistribution
538 in this complex topography (Fig. 11, S7, S8). This is particularly notable given the highly
539 variable magnitudes of accumulation over the five-year study and the contrasting climate
540 regions of these two glaciers (wet, warm maritime and cold, dry continental), with unique
541 storm paths, timing of annual accumulation, wind direction and wind direction
542 variability, and snow density. At both glaciers, the lowest interannual variability was
543 found away from locations with complex topography and elevated surface roughness,
544 such as crevassed zones, glacier margins, and areas near peaks and ridges.

545

546 In the most directly comparable study using repeat GPR surveys at Switzerland's
547 Findelgletscher, 86 % of the glacier area experienced less than 25 % range in relative
548 normalized accumulation over a three-year interval (Sold et al., 2016). As noted in
549 Section 3.4., we reported an absolute normalized range, whereas Sold et al. (2016)
550 reported a relative normalized range. Following their calculation, we found that 81 and
551 82 % of Wolverine and Gulkana's area experienced a relative normalized range less than
552 25 %. Collectively, our results add to the growing body of evidence (e.g., Deems et al.,
553 2008; Sturm and Wagner, 2010; Schirmer et al., 2011; Winstral and Marks, 2014)
554 suggesting 'time-stability' in the spatial distribution of snow in locations that span a
555 range of climate zones, topographic complexity, and relief. While the initial effort
556 required to constrain the spatial distribution over a given area can be significant, the
557 benefits of understanding the spatial distribution are substantial and long-lasting, and
558 have a wide range of applications.

559

560 **5.1.1 Elevation**

561 Elevation explained between 50 and 83 % of the observed SWE variability at Gulkana
562 and Wolverine, making it the most significant terrain parameter at both glaciers every
563 year (Fig. 4, 6). Steep winter SWE gradients characterized both glaciers throughout the
564 study period (115 – 440 mm 100 m⁻¹). Such gradients are comparable to previous results
565 for glaciers in the region (Pelto, 2008; Pelto et al., 2013; McGrath et al., 2015), but
566 exceed reported orographic precipitation gradients in other mountainous regions by a
567 factor of 2–3 (e.g., Anderson et al., 2014; Grünewald and Lehning, 2011). These steep
568 gradients are likely the result of physical processes beyond just orographic precipitation,
569 including storm systems that deliver snow at upper elevations and rain at lower elevations
570 (common at both Wolverine and Gulkana) and mid-winter ablation at lower elevations (at
571 Wolverine). These processes have also been shown to steepen observed SWE gradients
572 relative to orographic precipitation gradients in a mid-latitude seasonal snow watershed
573 (Anderson et al., 2014). Unfortunately, given that we solely sampled snow distribution at
574 the end of the accumulation season, the relative magnitude of each of these secondary
575 processes is not constrained.

576

577 Wolverine and Gulkana glaciers exhibited opposing SWE gradients at their highest
578 elevations, with Wolverine showing a sharp non-linear increase in SWE, while Gulkana
579 showed a gradual decrease. This non-linear increase was also noted at two maritime
580 glaciers (Scott and Valdez) in 2013 (McGrath et al., 2015), and perhaps reflects an
581 abundance of split precipitation phase storms in these warm coastal regions. The cause of
582 the observed reverse gradient at Gulkana may be the result of wind scouring at the
583 highest and most exposed sections of the glacier, or in part, a result of where we were
584 able to safely sample the glacier. For instance, in 2013, when we were able to access the
585 highest basin on the glacier, the SWE elevation gradient remained positive (Fig. 4).
586 Reductions in accumulated SWE at the highest elevations have also been observed at
587 Lemon Creek Glacier in southeast Alaska and Findel Glacier in Switzerland (Machguth
588 et al., 2006), presumably related to wind scouring at these exposed elevations.

589

590 **5.1.2. Wind redistribution**

591 Both statistical extrapolation approaches found terrain parameters Sb and curvature,
592 proxies for wind redistribution, to have the largest beta coefficients after elevation (Fig.
593 6, S9). The spatial pattern of SWE estimated by each model clearly reflects the dominant
594 influence of wind redistribution and elevation (Fig. 8), as areas of drift and scour are
595 apparent, especially at higher elevations. However, these terms do not fully capture the
596 redistribution process, as the model residuals (Fig. S1, S2) show sequential positive and
597 negative residuals associated with drift/scour zones. There are a number of reasons why
598 this might occur, including variable wind directions transporting snow (this is likely a
599 more significant issue at Gulkana, which experiences greater wind direction variability
600 (Fig. S6)), complex wind fields that are not well represented by a singular wind direction
601 (Dadic et al., 2010), changing surface topography (the glacier surface is dynamic over a
602 range of temporal scales, changing through both surface mass balance processes and ice
603 dynamics), and widely varying wind velocities. This is particularly relevant at Wolverine,
604 where wind speeds regularly gust over 30 m s^{-1} during winter storms, speeds that result in
605 variable length scales of redistribution that would not be captured by a fixed length scale
606 of redistribution. All of these factors influence the redistribution of snow and limit the
607 predictive ability of relatively simple proxies. Significant effort has gone into developing
608 physically-based snow-distribution models (e.g., Alpine3D and SnowModel), however,
609 high-resolution meteorological forcing data requirements generally limit the application
610 of these models in glacierized basins. Where such observations do exist, previous studies
611 have illuminated how the final distribution of snow is strongly correlated to the complex
612 wind field, including vertical (surface normal) winds (Dadic et al., 2010).

613

614 **5.1.3. Differences with non-glaciated terrain**

615 Although our GPR surveys did not regularly include non-glaciated regions of these
616 basins, a few key differences are worth noting. First, the length scales of variability on
617 and off the glacier were distinctly different, with shorter scales and greater absolute
618 variability (snow-free to $>5 \text{ m}$ in less than 10 m distance) off-glacier (Fig. S10). This
619 point has been clearly shown using airborne LiDAR in a glaciated catchment in the
620 Austrian Alps (Helfricht et al., 2014). The reduced variability on the glacier is largely due
621 to surface mass balance and ice flow processes that act to smooth the surface, leading to a

622 more spatially consistent surface topography, and therefore a more spatially consistent
623 SWE pattern. For this reason, establishing a SWE elevation gradient on a glacier is likely
624 much less prone to terrain-induced outliers compared to off-glacier sites, although the
625 relationship of this gradient to off-glacier gradients is generally unknown.

626

627 **5.2. Spatial differences between statistical models**

628 The two statistical extrapolation approaches yielded comparable large-scale spatial
629 distributions and glacier-wide averages, although there were some notable spatial
630 differences (Fig. 10). The systematic positive bias of the MVR approach over the
631 regression tree at Wolverine was due to three sectors of the glacier with both complex
632 terrain (i.e., icefalls) and large data gaps (typically locations that are not safe to access on
633 ground surveys). The difference in predicted SWE in these locations is likely due to how
634 the two statistical extrapolation approaches handle unsampled terrain parameter space.
635 The MVR extrapolates based on global linear trends, while the regression tree assigns
636 SWE from terrain that most closely resembles the under-sampled location. Anecdotally,
637 it appears that the MVR may overestimate SWE in some of these locations, which is most
638 evident in Wolverine's lower icefall, where bare ice is frequently exposed at the end of
639 the accumulation season (Fig. S11) in locations where the MVR predicted substantial
640 SWE. Likewise, the regression tree models could be underestimating SWE in these
641 regions, but in the absence of direct observations the errors are inherently unknown. The
642 regression tree model captures more short length scale variability while the MVR model
643 clarifies the larger trends. Consequently, smaller drifts and scours are captured well by
644 the regression tree model in areas where the terrain parameter space is well surveyed, but
645 the results become progressively less plausible as the terrain becomes distinctly different
646 from the sampled terrain parameter space. In contrast, the MVR model appears to give
647 more plausible results at larger spatial scales. This suggests that there is some theoretical
648 threshold where the regression tree is more appropriate if the terrain parameter space is
649 sampled sufficiently, but that for many glacier surveys the MVR model would be more
650 appropriate.

651

652 **5.3. Winter mass balance comparisons**

653 On average, all methods for estimating B_w were within ± 11 % of the six-method mean,
654 (Fig. 13). The agreement (as measured by the average percent difference from the mean)
655 between estimates was slightly better at Gulkana than Wolverine, likely reflecting the
656 overall lower spatial variability at Gulkana and the greater percentage of the glacier area
657 where b_w correlates well with the glacier-wide average (Fig. 11 e, f). At both glaciers, B_w
658 solutions based solely on elevation showed excellent agreement to the six-method mean,
659 suggesting that this simple approach is a viable means for measuring B_w on these glaciers.
660 The biggest differences occurred between the GPR-forced MVR model and the
661 glaciological site-index method, which we've shown is attributed to the upper stake (with
662 the greatest weight) underestimating the median SWE for that index zone (Fig. 14). The
663 upper stake location was established in 1966 at an elevation below the median elevation
664 of that index zone, which given the strong elevation control on SWE, is a likely reason
665 for the observed difference. At Gulkana, the relationship between the upper index site
666 and the GPR-forced MVR model is more variable in large part due to observed
667 differences in the accumulation between the main branch (containing the index site) and
668 the west branch of the glacier (containing additional stakes added in 2009). Such basin-
669 scale differences are likely present on many glaciers with complex geometry, and thus
670 illustrate potential uncertainties of using a small network of stakes to monitor the mass
671 balance of these glaciers. In the context of the MVR model, this manifests as a change in
672 sign in the eastness coefficient (which separates the branches in parameter space; Fig.
673 S4). Notably, in the two years where the site-index estimate was most negatively biased
674 at Gulkana (2015 and 2016), the glaciological profile method, relying on the more
675 extensive stake network (which includes stakes in the west branch of the glacier), yielded
676 B_w estimates within a few percent of the GPR-derived MVR estimate.

677

678 These GPR-derived B_w results have important implications for the cumulative
679 glaciological (stake-derived) mass balance time-series (currently only based on the site-
680 index method), which is calibrated with geodetic observations (details on the site-index
681 method and geodetic calibrations can be found in Van Beusekom et al., 2010 and O'Neel
682 et al., 2014). It is important to remember that the previous comparisons (e.g., Fig. 13)
683 were based on glaciological B_w values that have not had a geodetic calibration applied. At

684 Wolverine, the cumulative annual glaciological mass balance solutions are positively
685 biased compared to the geodetic mass balance solutions over decadal timescales,
686 requiring a negative calibration ($-0.43 \text{ m w.e. a}^{-1}$; O’Neel et al., 2014) to be applied to
687 the glaciological solutions. The source of this disagreement is some combination of the
688 stake-derived winter and summer balances being too positive relative to the geodetic
689 solution. On average, the GPR-derived B_w results were $\sim 0.4 \text{ m w.e.}$ more positive than the
690 site-index B_w results at Wolverine, which would further increase the glaciological-
691 geodetic solution difference and suggest that the stake-derived glaciological solutions are
692 underestimating ablation (B_s) by $\sim 0.8 \text{ m w.e. a}^{-1}$. Preliminary observations at Wolverine
693 using ablation wires show that some sectors of the glacier experience very high ablation
694 rates that are not captured by the stake network (e.g., crevassed zones through enhanced
695 shortwave solar radiation gain (e.g., Pfeffer and Bretherton, 1987; Cathles et al., 2011;
696 Colgan et al., 2016), and/or increased turbulent heat fluxes due to enhanced surface
697 roughness), and/or ice margins (through enhanced longwave radiation from nearby snow-
698 free land cover)). However, these results are not universal, as the assimilation of
699 distributed GPR observations at Findelgletcher significantly improved the comparison
700 between geodetic and modeled mass balance estimates (Sold et al., 2016), suggesting
701 multiple drivers of glaciologic-geodetic mismatch for long-term mass balance programs.

702 703 **5.3.1. Implications for stake placement**

704 Understanding the spatiotemporal distribution of SWE is useful for informing stake
705 placements and also for quantifying the uncertainty that interannual spatial variations in
706 SWE introduce to historic estimates of glacier-wide mass balance, particularly when
707 long-term mass balance programs rely on limited numbers of point observations (e.g.,
708 USGS and National Park Service glacier monitoring programs; O’Neel et al., 2014;
709 Burrows, 2014). Our winter balance results illustrate that stakes placed at the same
710 elevation are not directly comparable, and hence are not necessarily interchangeable in
711 the context of a multi-year mass balance record. Most locations on the glacier exhibit bias
712 from the average mass balance at that elevation and our results suggest interannual
713 consistency in this bias over sub-decadal time scales. As a result, constructing a balance

714 profile using a small number of inconsistently located stakes is likely to introduce large
715 relative errors from one year to the next.

716

717 Considering this finding, the placement of stakes to measure snow accumulation is
718 dependent on whether a single glacier-wide winter mass balance value (B_w) or a spatially
719 distributed SWE field is desired as a final product. For the former, a small number of
720 stakes can be distributed over the glacier hypsometry in areas where interannual
721 variability is low. Alternatively, if a distributed field is desired, a large number of stakes
722 can be widely distributed across the glacier, including areas where the interannual
723 variability is higher. In both cases it is important to have consistent locations from year to
724 year, although as the number of stakes increases significantly, this becomes less critical.

725

726 We assess the uncertainty that interannual variability in the spatial distribution of SWE
727 introduces to the historic index-method (March and Trabant, 1996) mass balance
728 solutions by first calculating the uncertainty, σ , contributed by each stake as:

$$729 \sigma_{stake} = \sigma_{model\ residuals} + (1 - R^2) \cdot u, \quad (3)$$

730 where $\sigma_{model\ residuals}$ is the standard deviation of MVR model residuals over all five
731 years within ± 30 meters of the index site, u is the mean b_w within ± 30 meters of the
732 index site, and R^2 is the coefficient of determination between b_w and B_w over the five-year
733 period (Fig. 11). The first term on the right hand side of Eq. 3 accounts for both the
734 spatial and temporal variability in the observed b_w as compared to the model, and the
735 second term accounts for the variability of the model as compared to B_w . The glacier-
736 wide uncertainty from interannual variability is then:

$$737 \text{Glacier } \sigma = \sqrt{\sum_{all\ stakes} (\sigma_{stake} \cdot w_{stake})^2}, \quad (4)$$

738 where w_{stake} is the weight function from the site-index method (which depends on stake
739 location and glacier hypsometry). By this assessment, interannual variability in the spatial
740 distribution of SWE at stake locations introduced minor uncertainty, on the order of 0.11
741 m w.e. at both glaciers (4 % and 10 % of B_w at Wolverine and Gulkana, respectively).

742 This suggests that the original stake network design at the benchmark glaciers does
743 remarkably well at capturing the interannual variability in glacier-wide winter balance.

744 The greatest interannual variability at each glacier is found at the lowest stake sites, but

745 because b_w and the stake weights are both quite low at these sites, they contribute only
746 modestly to the overall uncertainty. Instead, the middle and upper elevation stakes
747 contribute the greatest amount to the glacier-wide uncertainty.

748

749 **6. Conclusions**

750 We collected spatially extensive GPR observations at two glaciers in Alaska for five
751 consecutive winters to quantify the spatiotemporal distribution of SWE. We found good
752 agreement of glacier-average winter balances, B_w , among the four different approaches
753 used to extrapolate GPR point measurements of SWE across the glacier hypsometry.
754 Extrapolations relying only on elevation (i.e., a simple balance profile) produced B_w
755 estimates similar to the more complicated statistical models, suggesting that this is an
756 appropriate method for quantifying glacier-wide winter balances at these glaciers. The
757 more complicated approaches, which allow SWE to vary across a range of terrain-
758 parameters based on DEMs, show a high degree of temporal stability in the pattern of
759 accumulation at both glaciers, as ~85 % of the area on both glaciers experienced less than
760 25 % normalized absolute variability over the five-year interval. Elevation and the
761 parameters related to wind redistribution had the most explanatory power, and were
762 temporally consistent at each site. The choice between MVR and regression tree models
763 should depend on both the range in terrain parameter space that exists on the glacier,
764 along with how well that space is surveyed.

765

766 In total, six different methods (four based on GPR measurements and two based on stake
767 measurements) for estimating the glacier-wide average agreed within ± 11 %. The site-
768 index glaciological B_w estimates were negatively biased compared to all other estimates,
769 particularly when the upper-elevation stake significantly underestimated SWE in that
770 index zone. In contrast, the profile glaciological approach, using a more extensive stake
771 network, showed better agreement with the other approaches, highlighting the benefits of
772 using a more extensive stake network.

773

774 We found the spatial patterns of snow accumulation to be temporally stable on these
775 glaciers, which is consistent with a growing body of literature documenting similar

776 consistency in a wide variety of environments. The long-term stake locations experienced
777 low interannual variability in normalized SWE, meaning that stake measurements tracked
778 the interannual variability in SWE, rather than interannual variability in spatial patterns.
779 The uncertainty associated with interannual spatial variability is only 4–10 % of the
780 glacier-wide B_w at each glacier. Thus, our findings support the concept that sparse stake
781 networks can be effectively used to measure interannual variability in winter balance on
782 glaciers.

783

784 *Data Availability.* The GPR and associated observational data used in this study can be
785 accessed on the USGS Glaciers and Climate Project website
786 (<https://doi.org/10.5066/F7M043G7>). The Benchmark Glacier mass balance input and
787 output can be accessed at: <https://doi.org/10.5066/F7HD7SRF> (O’Neel et al., 2018). The
788 Gulkana DEM is available from the ArcticDEM project website
789 (<https://www.pgc.umn.edu/data/arcticdem/>) and the Wolverine DEM is available at
790 <ftp://bering.gps.alaska.edu/pub/chris/wolverine/>. A generalized version of the SWE
791 extrapolation code is available at: [https://github.com/danielmcgrathCSU/Snow-](https://github.com/danielmcgrathCSU/Snow-Distribution)
792 [Distribution](https://github.com/danielmcgrathCSU/Snow-Distribution).

793

794 *Author Contributions.* SO, DM, LS, and HPM designed the study. DM performed the
795 analyses and wrote the manuscript. LS contributed to the design and implementation of
796 the analyses, and CM, SC, and EHB contributed specific components of the analyses. All
797 authors provided feedback and edited the manuscript.

798

799 *Competing Interests.* The authors declare that they have no conflict of interest.

800

801 *Acknowledgments.* This work was funded by the U.S. Geological Survey Land Change
802 Science Program, USGS Alaska Climate Adaptation Science Center, and DOI/USGS
803 award G17AC00438 to DM. Any use of trade, firm, or product names is for descriptive
804 purposes only and does not imply endorsement by the U.S. Government. We
805 acknowledge the Polar Geospatial Center (NSF-OPP awards 1043681, 1559691, and
806 1542736) for the Gulkana DEM. We thank Caitlyn Florentine, Jeremy Littell, Mauri

807 Pelto, and an anonymous reviewer for their thoughtful feedback that improved the
808 manuscript.

809

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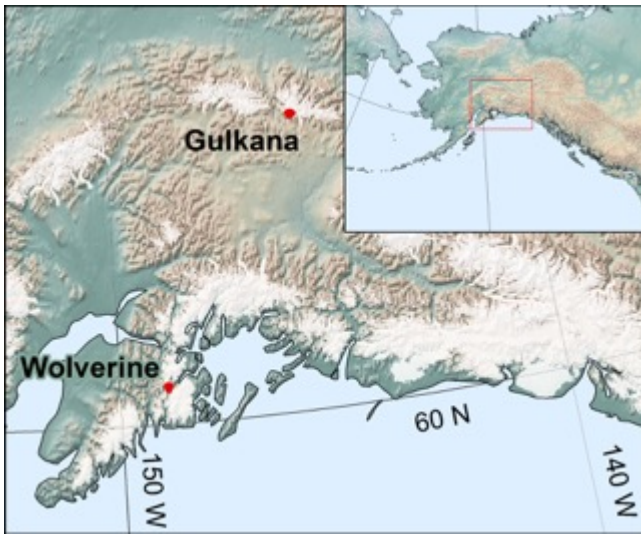
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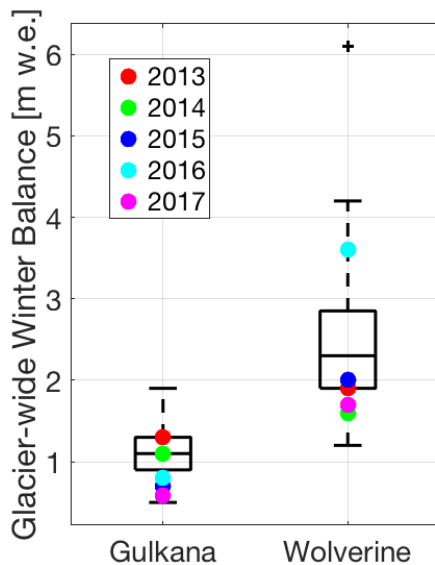
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1050 Figure 1. Map of southern Alaska with study glaciers marked by red outline. All glaciers
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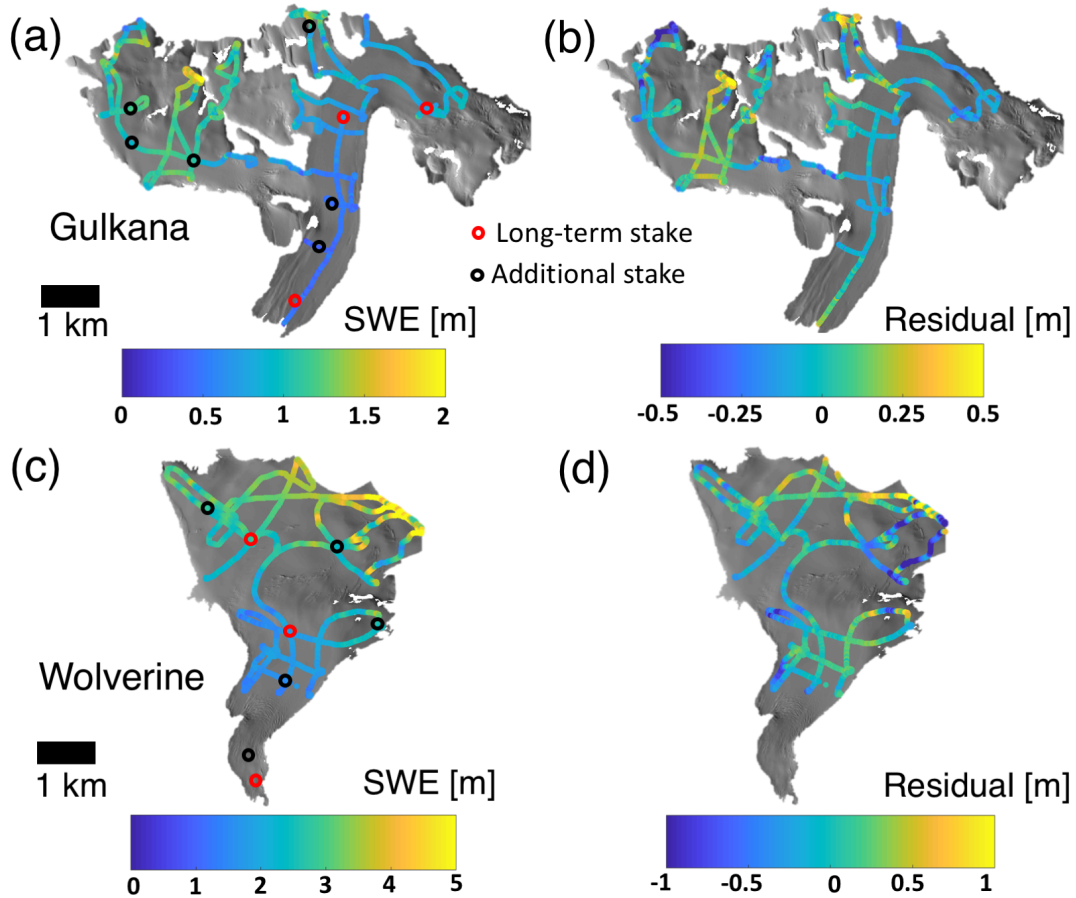
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1057 Figure 2. Boxplots of glacier-wide winter balance for Gulkana and Wolverine glaciers
1058 between 1966 and 2017. Years corresponding to GPR surveys are shown with colored
1059 markers. These values have not been adjusted by the geodetic calibration (see O’Neel et
1060 al., 2014).



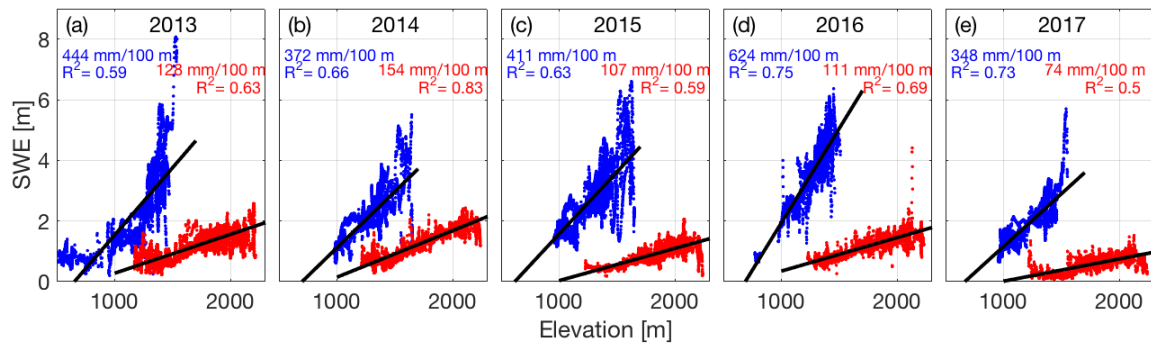
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1065 Figure 3. GPR surveys from 2015 at Gulkana (a) and Wolverine (c) glaciers and MVR
 1066 model residuals (b, d).



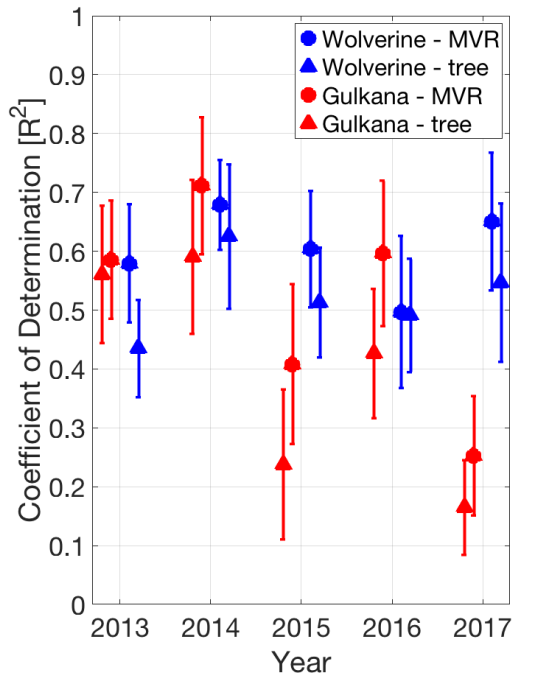
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1070 Figure 4. SWE from GPR surveys as a function of elevation, along with least squares
 1071 regression slope and coefficient of determination for each year of the study period.
 1072 Wolverine is plotted in blue, Gulkana in red.



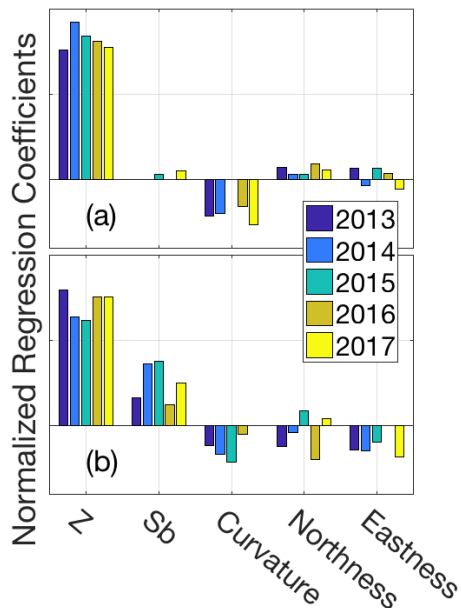
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1075 Figure 5. Median and standard deviation (error bars) of coefficient of determination
 1076 (from 100 model runs) for both extrapolation approaches (circles are MVR, triangles are
 1077 regression tree) developed on training datasets and applied to test datasets. Symbols and
 1078 error bars are offset from year for clarity.



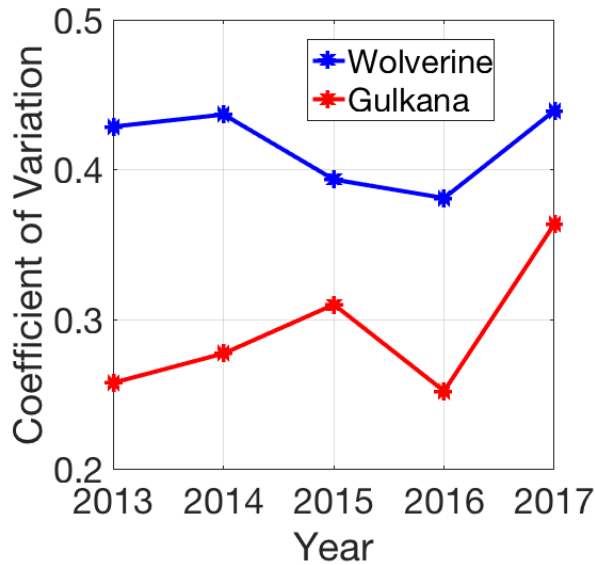
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1097 Figure 6. Terrain parameter beta coefficients for (a) Gulkana and (b) Wolverine for
 1098 multivariable linear regression for each year of the study interval.



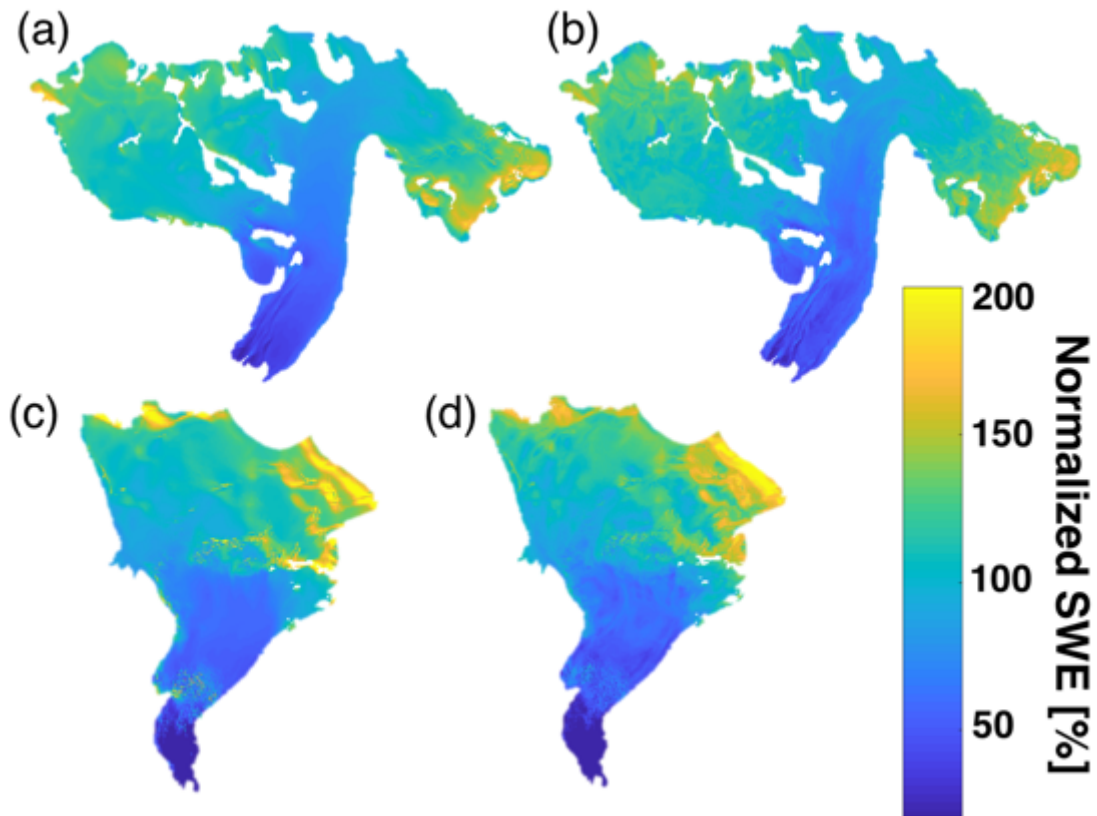
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1100 Figure 7. Spatial variability in snow accumulation across the glacier quantified by the
1101 coefficient of variation (standard deviation/mean) for each glacier across the five-year
1102 interval based on MVR model output.
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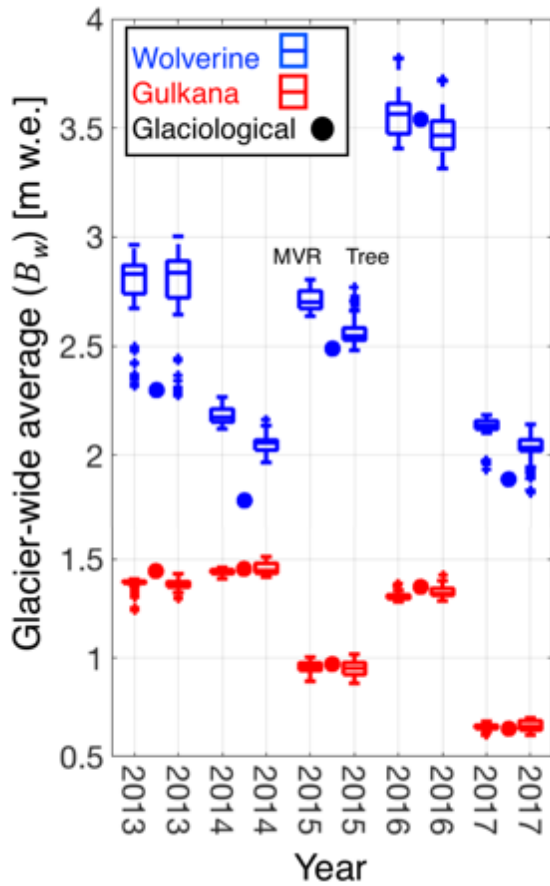
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1106 Figure 8. Five-year mean of normalized distributed SWE for Gulkana (a,b) and
1107 Wolverine (c,d) for multivariable regression (a,c) and regression tree (b,d).



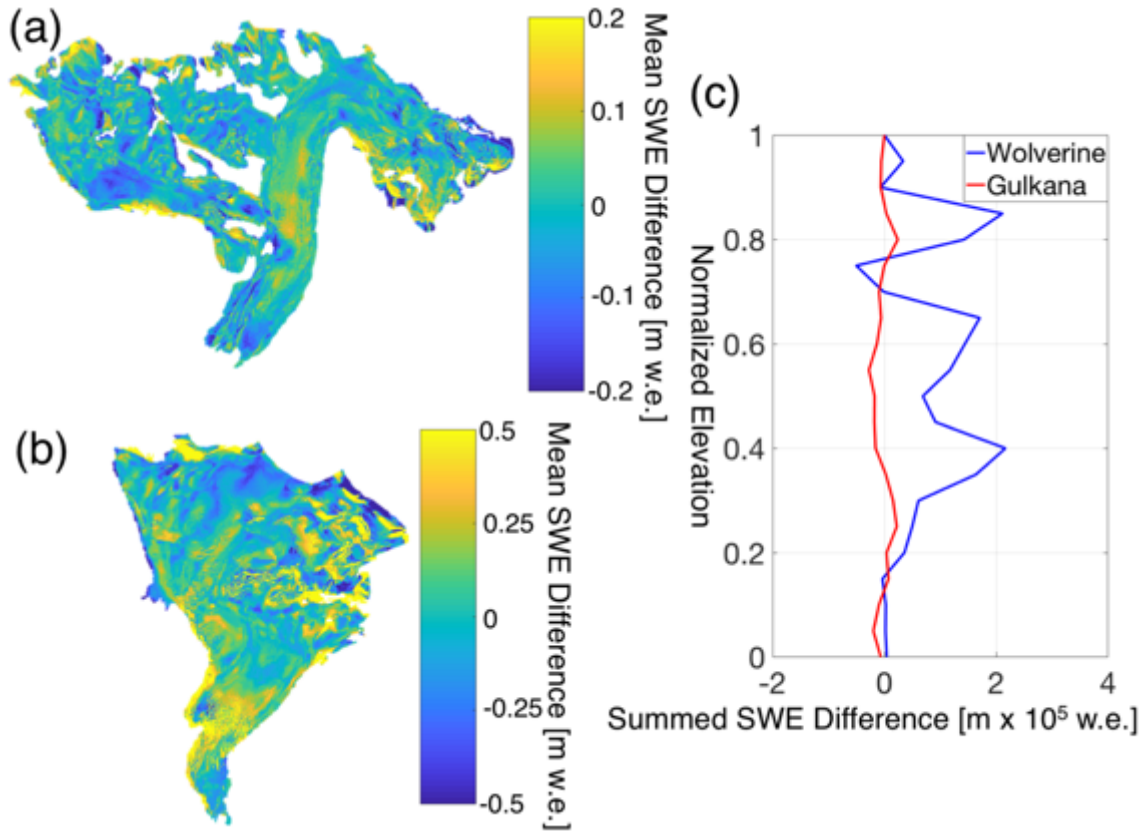
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1109 Figure 9. Comparing statistical models for GPR-derived glacier-wide winter balances for
 1110 both Wolverine (blue) and Gulkana (red) glaciers. For each year and each glacier, two
 1111 boxplots are shown. The first shows multivariable regression model (MVR) output and
 1112 the second shows regression tree output (tree). The B_w estimate from the glaciological
 1113 profile method is shown for each year and glacier as the filled circle.
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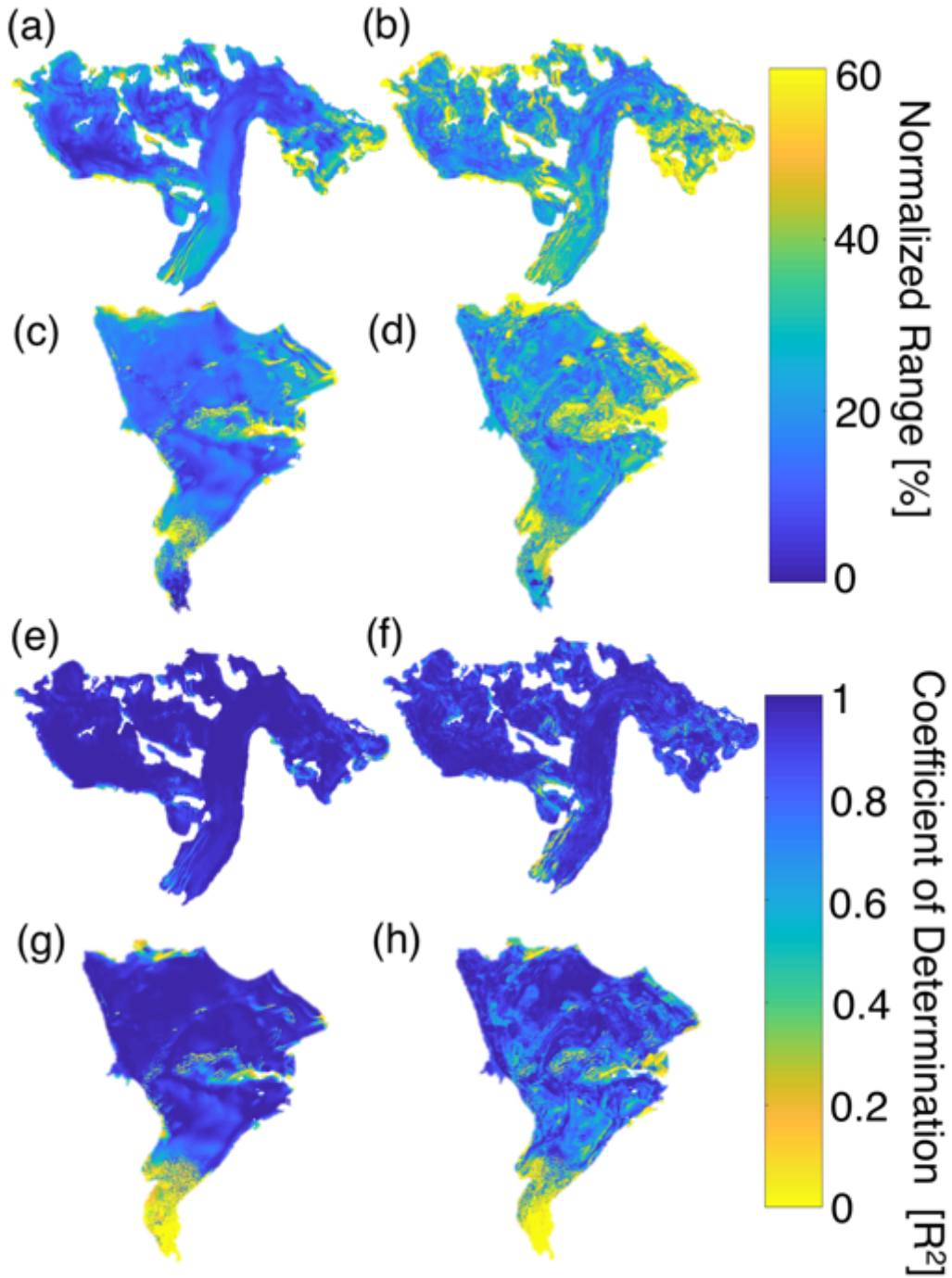
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1126 Figure 10. SWE differences between statistical models for Gulkana (a) and Wolverine
1127 (b) calculated by differencing the regression tree five-year mean SWE from the
1128 multivariable regression (MVR) five-year mean SWE. Yellow colors indicate regions
1129 where MVR yields more SWE than decision tree and blue colors indicate the opposite.
1130 Note different magnitude colorbar scales. c) Summed SWE difference between methods
1131 in bins of 0.05 normalized elevation values.



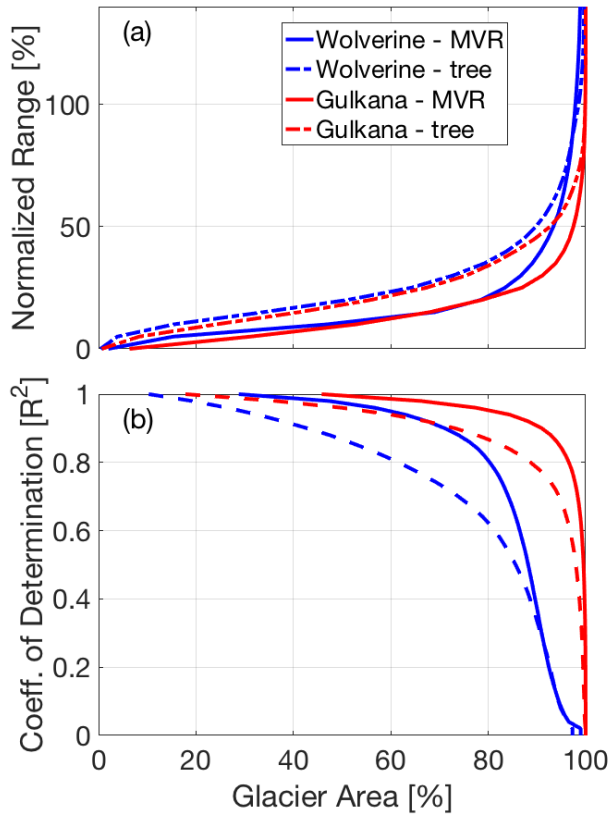
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1146 Figure 11. Interannual variability of the SWE accumulation field from 2013–2017,
1147 quantified via normalized range (a-d) and R^2 (e-h) approach for median distributed fields
1148 from the multivariable regression (left column) and regression tree (right column)
1149 statistical models.



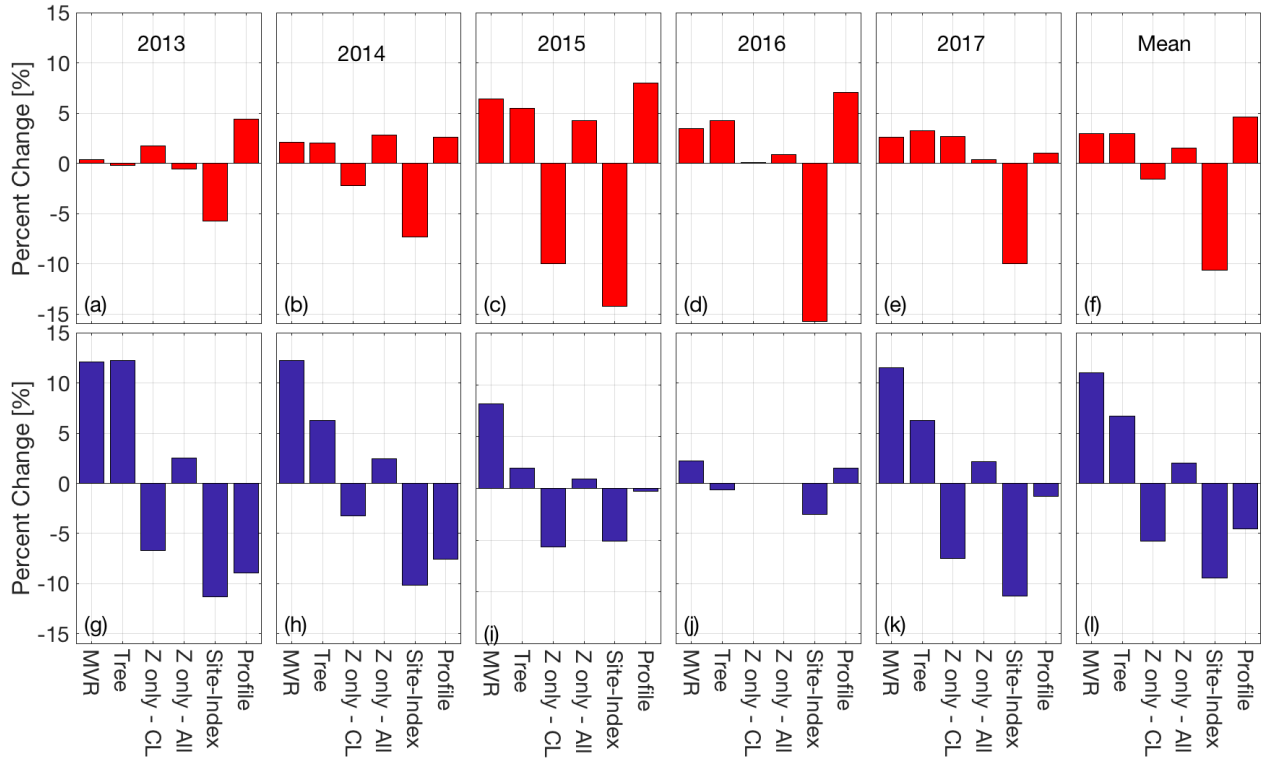
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1152 Figure 12. Interannual variability of the SWE accumulation pattern as a function of
1153 cumulative glacier area, shown as (a) normalized range and (b) and R^2 . Solid lines are for
1154 multivariable regression (MVR) and dashed lines are regression tree.
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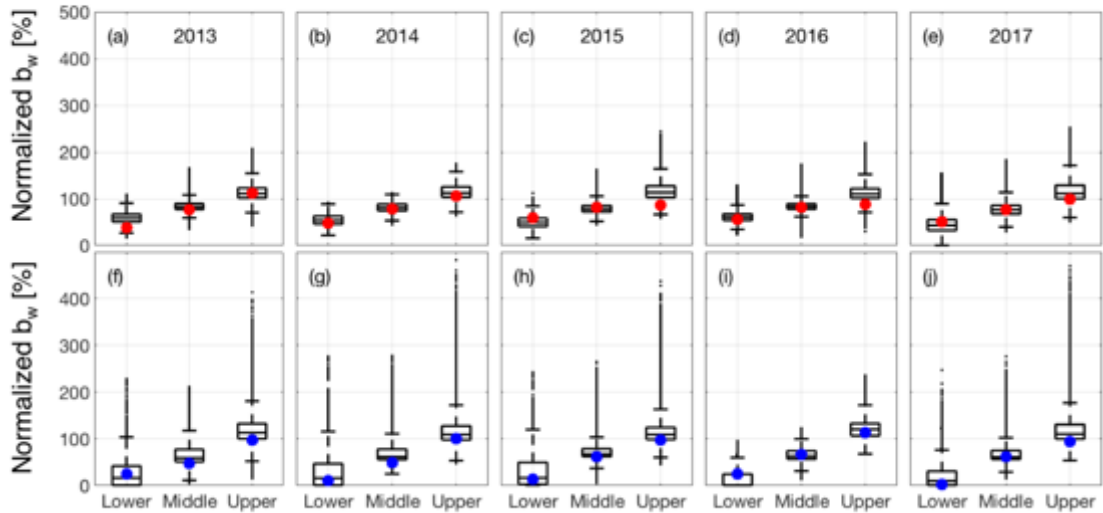
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1171 Figure 13. Percent deviation for each estimate from the six-method mean of B_w .
 1172 Individual years for Gulkana Glacier are shown in panels a-e with the five-year mean
 1173 shown in f. Individual years for Wolverine Glacier are shown in panels g-k, with the five-
 1174 year mean shown in l.



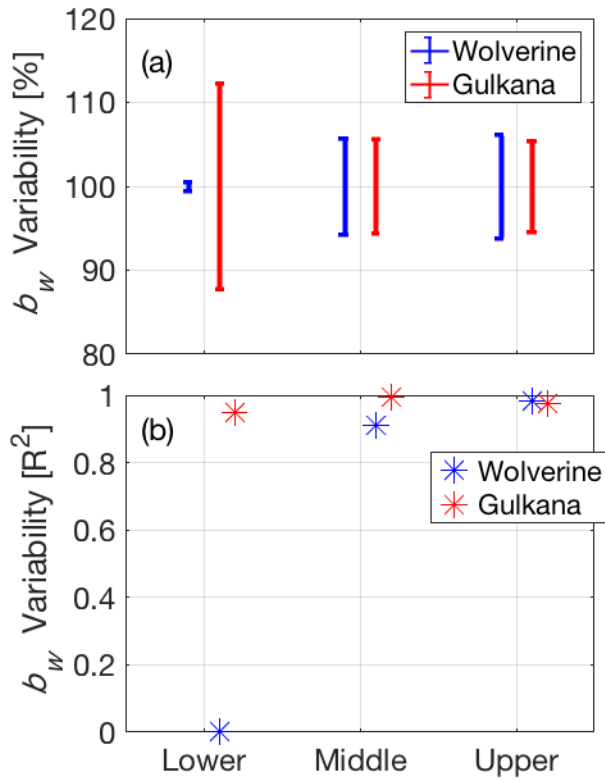
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1191 Figure 14. Spatial variability in snow accumulation for individual years (2013-2017) by
 1192 elevation (lower, middle, upper) compared to stake measurements. Box plot of all
 1193 distributed SWE values (from multivariable regression) for each index zone of the glacier
 1194 for Gulkana (a-e) and Wolverine (f-j) for 2013-2017. The filled circles are the respective
 1195 stake observation for that index zone. SWE is expressed as a percentage of the glacier-
 1196 wide average, B_w , for that year and glacier.



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1217 Figure 15. Interannual variability in the spatial pattern of snow accumulation at long-term
 1218 mass balance stake locations for Wolverine and Gulkana glaciers using a) normalized b_w
 1219 range and b) coefficient of determination (from Figure 11; MVR model).
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